Multioutput Gaussian Processes

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Parametric Models are a Bottleneck

Constructing Covariance

GP Limitations

Kalman Filter



Parametric Models are a Bottleneck

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- This work takes us from parametric to non-parametric.
- The limit implies infinite dimensional **w**.
- Gaussian processes are generally non-parametric: combine data with covariance function to get model.
- This representation *cannot* be summarized by a parameter vector of a fixed size.

- Parametric models have a representation that does not respond to increasing training set size.
- Bayesian posterior distributions over parameters contain the information about the training data.
 - ► Use Bayes' rule from training data, *p*(**w**|**y**, **X**),
 - Make predictions on test data

$$p(y_*|\mathbf{X}_*, \mathbf{y}, \mathbf{X}) = \int p(y_*|\mathbf{w}, \mathbf{X}_*) p(\mathbf{w}|\mathbf{y}, \mathbf{X}) d\mathbf{w}).$$

- w becomes a bottleneck for information about the training set to pass to the test set.
- Solution: increase *m* so that the bottleneck is so large that it no longer presents a problem.
- How big is big enough for *m*? Non-parametrics says $m \rightarrow \infty$.

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- These are known as degenerate covariance matrices.
- Their rank is at most *m*, non-parametric models have full rank covariance matrices.
- Most well known is the "linear kernel", $k(\mathbf{x}_i, \mathbf{x}_j) = \mathbf{x}_i^{\top} \mathbf{x}_j$.

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- Parametric is a special case when conditional prediction can be summarized in a *fixed* number of parameters.
- Complexity of parametric model remains fixed regardless of the size of our training data set.
- For a non-parametric model the required number of parameters grows with the size of the training data.

Covariance Functions and Mercer Kernels

Mercer Kernels and Covariance Functions are similar.

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Covariance Functions and Mercer Kernels

- Mercer Kernels and Covariance Functions are similar.
- the kernel perspective does not make a probabilistic interpretation of the covariance function.
- Algorithms can be simpler, but probabilistic interpretation is crucial for kernel parameter optimization.



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Constructing Covariance Functions

Sum of two covariances is also a covariance function.

$$k(\mathbf{x}, \mathbf{x}') = k_1(\mathbf{x}, \mathbf{x}') + k_2(\mathbf{x}, \mathbf{x}')$$

Constructing Covariance Functions

Product of two covariances is also a covariance function.

 $k(\mathbf{x},\mathbf{x}')=k_1(\mathbf{x},\mathbf{x}')k_2(\mathbf{x},\mathbf{x}')$

Multiply by Deterministic Function

- If $f(\mathbf{x})$ is a Gaussian process.
- $g(\mathbf{x})$ is a deterministic function.
- $h(\mathbf{x}) = f(\mathbf{x})g(\mathbf{x})$
- Then

$$k_h(\mathbf{x}, \mathbf{x}') = g(\mathbf{x})k_f(\mathbf{x}, \mathbf{x}')g(\mathbf{x}')$$

where k_h is covariance for $h(\cdot)$ and k_f is covariance for $f(\cdot)$.

MLP Covariance Function

$$k(\mathbf{x}, \mathbf{x}') = \alpha \operatorname{asin}\left(\frac{w\mathbf{x}^{\top}\mathbf{x}' + b}{\sqrt{w\mathbf{x}^{\top}\mathbf{x} + b + 1}\sqrt{w\mathbf{x}'^{\top}\mathbf{x}' + b + 1}}\right)$$

 Based on infinite neural network model.

$$w = 40$$
$$b = 4$$



MLP Covariance Function

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 Based on infinite neural network model.

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Linear Covariance Function

$$k(\mathbf{x}, \mathbf{x}') = \alpha \mathbf{x}^\top \mathbf{x}'$$



$$\alpha = 1$$



Linear Covariance Function

$$k\left(\mathbf{x},\mathbf{x}'\right) = \alpha \mathbf{x}^{\top} \mathbf{x}'$$

Bayesian linear regression.

$$\alpha = 1$$



















Gaussian Noise

Gaussian noise model,

$$p(y_i|f_i) = \mathcal{N}(y_i|f_i,\sigma^2)$$

where σ^2 is the variance of the noise.

• Equivalent to a covariance function of the form

$$k(\mathbf{x}_i, \mathbf{x}_j) = \delta_{i,j} \sigma^2$$

where $\delta_{i,j}$ is the Kronecker delta function.

 Additive nature of Gaussians means we can simply add this term to existing covariance matrices.

Gaussian Process Regression



Figure: Examples include WiFi localization, C14 callibration curve.
















Can we determine covariance parameters from the data?

$$\mathcal{N}(\mathbf{y}|\mathbf{0},\mathbf{K}) = \frac{1}{(2\pi)^{\frac{n}{2}}|\mathbf{K}|} \exp\left(-\frac{\mathbf{y}^{\mathsf{T}}\mathbf{K}^{-1}\mathbf{y}}{2}\right)$$

The parameters are *inside* the covariance function (matrix).

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The parameters are *inside* the covariance function (matrix).

Can we determine covariance parameters from the data?

$$\log \mathcal{N}(\mathbf{y}|\mathbf{0}, \mathbf{K}) = -\frac{1}{2} \log |\mathbf{K}| - \frac{\mathbf{y}^{\mathsf{T}} \mathbf{K}^{-1} \mathbf{y}}{2} - \frac{n}{2} \log 2\pi$$

The parameters are *inside* the covariance function (matrix).

Can we determine covariance parameters from the data?

$$E(\boldsymbol{\theta}) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^{\top} \mathbf{K}^{-1} \mathbf{y}}{2}$$

The parameters are *inside* the covariance function (matrix).

Eigendecomposition of Covariance

A useful decomposition for understanding the objective function.

 $\mathbf{K} = \mathbf{R} \mathbf{\Lambda}^2 \mathbf{R}^\top$



Diagonal of Λ represents distance along axes. **R** gives a rotation of these axes.

where Λ is a *diagonal* matrix and $\mathbf{R}^{\top}\mathbf{R} = \mathbf{I}$.



















$$|\mathbf{\Lambda}| = \lambda_1 \lambda_2 \lambda_3$$





 $|\mathbf{R}\mathbf{\Lambda}| = \lambda_1 \lambda_2$

Data Fit: $\frac{\mathbf{y}^{-1}\mathbf{K}^{-1}\mathbf{y}}{2}$



 y_1

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Gene Expression Example



Data from Della Gatta et al. (2008). Application from Kalaitzis and Lawrence (2011).


Contour plot of Gaussian



Optima: length scale of 1.2221 and \log_{10} SNR of 1.9654 log likelihood is -0.22317.



Optima: length scale of 1.5162 and \log_{10} SNR of 0.21306 log likelihood is -0.23604.



Optima: length scale of 2.9886 and \log_{10} SNR of -4.506 log likelihood is -2.1056.



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- ► Inference is O(n³) due to matrix inverse (in practice use Cholesky).
- Gaussian processes don't deal well with discontinuities (financial crises, phosphorylation, collisions, edges in images).
- Widely used exponentiated quadratic covariance (RBF) can be too smooth in practice (but there are many alternatives!!).

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Simple Markov Chain

- Assume 1-d latent state, a vector over time, $\mathbf{x} = [x_1 \dots x_T]$.
- Markov property,

$$x_{i} = x_{i-1} + \epsilon_{i},$$

$$\epsilon_{i} \sim \mathcal{N}(0, \alpha)$$

$$\Rightarrow x_{i} \sim \mathcal{N}(x_{i-1}, \alpha)$$

Initial state,

 $x_0 \sim \mathcal{N}(0, \alpha_0)$

- If $x_0 \sim \mathcal{N}(0, \alpha)$ we have a Markov chain for the latent states.
- Markov chain it is specified by an initial distribution (Gaussian) and a transition distribution (Gaussian).

=



















Multivariate Gaussian Properties: Reminder

If $\mathbf{z} \sim \mathcal{N}(\mu, \mathbf{C})$ and $\mathbf{x} = \mathbf{W}\mathbf{z} + \mathbf{b}$ then $\mathbf{x} \sim \mathcal{N}(\mathbf{W}\mu + \mathbf{b}, \mathbf{W}\mathbf{C}\mathbf{W}^{\mathsf{T}})$

Multivariate Gaussian Properties: Reminder





 $x_1 = \epsilon_1$



 $x_2 = \epsilon_1 + \epsilon_2$



 $x_3 = \epsilon_1 + \epsilon_2 + \epsilon_3$



 $x_4 = \epsilon_1 + \epsilon_2 + \epsilon_3 + \epsilon_4$



 $x_5 = \epsilon_1 + \epsilon_2 + \epsilon_3 + \epsilon_4 + \epsilon_5$

 $\mathbf{x} = \mathbf{L}_1 \times \boldsymbol{\epsilon}$

- Since x is linearly related to *e* we know x is a Gaussian process.
- Trick: we only need to compute the mean and covariance of x to determine that Gaussian.

$x = L_1 \epsilon$

$\langle x \rangle = \langle L_1 \varepsilon \rangle$

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$\langle x\rangle = L_1 \langle \varepsilon \rangle$

$\boldsymbol{\epsilon} \sim \mathcal{N}\left(\mathbf{0}, \alpha \mathbf{I}\right)$

$\langle x\rangle = L_1 0$

$\langle x \rangle = 0$

$\begin{aligned} \mathbf{x}\mathbf{x}^{\top} &= \mathbf{L}_{1}\boldsymbol{\epsilon}\boldsymbol{\epsilon}^{\top}\mathbf{L}_{1}^{\top} \\ \mathbf{x}^{\top} &= \boldsymbol{\epsilon}^{\top}\mathbf{L}^{\top} \end{aligned}$

 $\langle \mathbf{x}\mathbf{x}^{\top}\rangle = \left\langle \mathbf{L}_{1}\boldsymbol{\epsilon}\boldsymbol{\epsilon}^{\top}\mathbf{L}_{1}^{\top}\right\rangle$

$\langle \mathbf{x}\mathbf{x}^{\top} \rangle = \mathbf{L}_1 \langle \boldsymbol{\epsilon}\boldsymbol{\epsilon}^{\top} \rangle \mathbf{L}_1^{\top}$

$\langle \mathbf{x}\mathbf{x}^{\mathsf{T}} \rangle = \mathbf{L}_1 \langle \boldsymbol{\epsilon}\boldsymbol{\epsilon}^{\mathsf{T}} \rangle \mathbf{L}_1^{\mathsf{T}}$

 $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{I})$
Latent Process Covariance

$\langle \mathbf{x}\mathbf{x}^{\top} \rangle = \alpha \mathbf{L}_{1}\mathbf{L}_{1}^{\top}$

$\mathbf{x} = \mathbf{L}_1 \boldsymbol{\epsilon}$

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$\mathbf{x} = \mathbf{L}_{1}\boldsymbol{\epsilon}$ $\boldsymbol{\epsilon} \sim \mathcal{N}\left(\mathbf{0}, \alpha \mathbf{I}\right)$ \Longrightarrow $\mathbf{x} \sim \mathcal{N}\left(\mathbf{0}, \alpha \mathbf{L}_{1}\mathbf{L}_{1}^{\top}\right)$

- Make the variance dependent on time interval.
- Assume variance grows *linearly* with time.
- Justification: sum of two Gaussian distributed random variables is distributed as Gaussian with sum of variances.
- If variable's movement is additive over time (as described) variance scales linearly with time.

► Given $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{I}) \Longrightarrow \boldsymbol{\epsilon} \sim \mathcal{N}\left(\mathbf{0}, \alpha \mathbf{L}_{1}\mathbf{L}_{1}^{\top}\right).$ Then $\boldsymbol{\epsilon} \sim \mathcal{N}\left(\mathbf{0}, \Delta t \alpha \mathbf{I}\right) \Longrightarrow \boldsymbol{\epsilon} \sim \mathcal{N}\left(\mathbf{0}, \Delta t \alpha \mathbf{L}_{1}\mathbf{L}_{1}^{\top}\right).$

where Δt is the time interval between observations.

$$\boldsymbol{\epsilon} \sim \mathcal{N}\left(0, \alpha \Delta t \mathbf{I}\right), \quad \mathbf{x} \sim \mathcal{N}\left(0, \alpha \Delta t \mathbf{L}_{1} \mathbf{L}_{1}^{\top}\right)$$

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 $\mathbf{K} = \alpha \Delta t \mathbf{L}_{\mathbf{1}} \mathbf{L}_{\mathbf{1}}^{\top}$

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$$\mathbf{K} = \alpha \Delta t \mathbf{L}_{\mathbf{1}} \mathbf{L}_{\mathbf{1}}^{\mathsf{T}}$$

$$k_{i,j} = \alpha \Delta t \mathbf{l}_{:,i}^{\top} \mathbf{l}_{:,j}$$

where $\mathbf{l}_{:,k}$ is a vector from the *k*th row of \mathbf{L}_1 : the first *k* elements are one, the next T - k are zero.

$$\boldsymbol{\epsilon} \sim \mathcal{N}\left(0, \alpha \Delta t \mathbf{I}\right), \quad \mathbf{x} \sim \mathcal{N}\left(0, \alpha \Delta t \mathbf{L}_{1} \mathbf{L}_{1}^{\top}\right)$$

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 $k_{i,j} = \alpha \Delta t \min(i, j)$ define $\Delta ti = t_i$ so $k_{i,j} = \alpha \min(t_i, t_j) = k(t_i, t_j)$

Where did this covariance matrix come from?

Markov Process

$$k(t,t') = \alpha \min(t,t')$$

 Covariance matrix is built using the *inputs* to the function *t*.



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Markov Process

Visualization of inverse covariance (precision).

- Precision matrix is sparse: only neighbours in matrix are non-zero.
- This reflects *conditional* independencies in data.
- In this case Markov structure.



Where did this covariance matrix come from?

Exponentiated Quadratic Kernel Function (RBF, Squared Exponential, Gaussian)

$$k(\mathbf{x}, \mathbf{x}') = \alpha \exp\left(-\frac{\|\mathbf{x} - \mathbf{x}'\|_2^2}{2\ell^2}\right)$$

- Covariance matrix is built using the *inputs* to the function x.
- For the example above it was based on Euclidean distance.
- The covariance function is also know as a kernel.



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Exponentiated Quadratic

Visualization of inverse covariance (precision).

- Precision matrix is not sparse.
- Each point is dependent on all the others.
- In this case non-Markovian.



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Markov Process

Visualization of inverse covariance (precision).

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Simple Kalman Filter I

• We have state vector $\mathbf{X} = [\mathbf{x}_1 \dots \mathbf{x}_q] \in \mathbb{R}^{T \times q}$ and if each state evolves independently we have

$$p(\mathbf{X}) = \prod_{i=1}^{q} p(\mathbf{x}_{:,i})$$
$$p(\mathbf{x}_{:,i}) = \mathcal{N}(\mathbf{x}_{:,i}|\mathbf{0}, \mathbf{K}).$$

• We want to obtain outputs through:

$$\mathbf{y}_{i,:} = \mathbf{W}\mathbf{x}_{i,:}$$

Stacking and Kronecker Products I

Represent with a 'stacked' system:

$$p(\mathbf{x}) = \mathcal{N}\left(\mathbf{x}|\mathbf{0}, \mathbf{I} \otimes \mathbf{K}\right)$$

where the stacking is placing each column of **X** one on top of another as

$$\mathbf{x} = \begin{bmatrix} \mathbf{x}_{:,1} \\ \mathbf{x}_{:,2} \\ \vdots \\ \mathbf{x}_{:,q} \end{bmatrix}$$

Kronecker Product



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Column Stacking















Can also stack each row of **X** to form column vector:

$$\mathbf{x} = \begin{bmatrix} \mathbf{x}_{1,:} \\ \mathbf{x}_{2,:} \\ \vdots \\ \mathbf{x}_{T,:} \end{bmatrix}$$

 $p(\mathbf{x}) = \mathcal{N}\left(\mathbf{x}|\mathbf{0}, \mathbf{K} \otimes \mathbf{I}\right)$

Row Stacking












The observations are related to the latent points by a linear mapping matrix,

$$\mathbf{y}_{i,:} = \mathbf{W}\mathbf{x}_{i,:} + \boldsymbol{\epsilon}_{i,:}$$
$$\boldsymbol{\epsilon} \sim \mathcal{N}\left(0, \sigma^2 \mathbf{I}\right)$$

Mapping from Latent Process to Observed



This leads to a covariance of the form

 $(\mathbf{I} \otimes \mathbf{W})(\mathbf{K} \otimes \mathbf{I})(\mathbf{I} \otimes \mathbf{W}^{\top}) + \mathbf{I}\sigma^{2}$ Using $(\mathbf{A} \otimes \mathbf{B})(\mathbf{C} \otimes \mathbf{D}) = \mathbf{A}\mathbf{C} \otimes \mathbf{B}\mathbf{D}$ This leads to $\mathbf{K} \otimes \mathbf{W}\mathbf{W}^{\top} + \mathbf{I}\sigma^{2}$

or

$$\mathbf{y} \sim \mathcal{N}\left(\mathbf{0}, \mathbf{W}\mathbf{W}^\top \otimes \mathbf{K} + \mathbf{I}\sigma^2\right)$$

Kernels for Vector Valued Outputs: A Review

Foundations and Trends[®] in Machine Learning Vol. 4, No. 3 (2011) 195–266 © 2012 M. A. Álvarez, L. Rosasco and N. D. Lawrence DOI: 10.1561/2200000036



Kernels for Vector-Valued Functions: A Review

By Mauricio A. Álvarez, Lorenzo Rosasco and Neil D. Lawrence This Kronecker structure leads to several published models.

$$(\mathbf{K}(\mathbf{x},\mathbf{x}'))_{d,d'}=k(\mathbf{x},\mathbf{x}')k_T(d,d'),$$

where *k* has **x** and k_T has *n* as inputs.

- Can think of multiple output covariance functions as covariances with augmented input.
- Alongside **x** we also input the *d* associated with the *output* of interest.

► Taking B = WW^T we have a matrix expression across outputs.

$$\mathbf{K}(\mathbf{x},\mathbf{x}')=k(\mathbf{x},\mathbf{x}')\mathbf{B},$$

where **B** is a $p \times p$ symmetric and positive semi-definite matrix.

- **B** is called the *coregionalization* matrix.
- We call this class of covariance functions *separable* due to their product structure.

Sum of Separable Covariance Functions

In the same spirit a more general class of kernels is given by

$$\mathbf{K}(\mathbf{x},\mathbf{x}') = \sum_{j=1}^{q} k_j(\mathbf{x},\mathbf{x}')\mathbf{B}_j.$$

This can also be written as

$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \sum_{j=1}^{q} \mathbf{B}_{j} \otimes k_{j}(\mathbf{X},\mathbf{X}),$$

- This is like several Kalman filter-type models added together, but each one with a different set of latent functions.
- We call this class of kernels sum of separable kernels (SoS kernels).

- Use of GPs in Geostatistics is called kriging.
- These multi-output GPs pioneered in geostatistics: prediction over vector-valued output data is known as *cokriging*.
- The model in geostatistics is known as the *linear model of* coregionalization (LMC, Journel and Huijbregts (1978); Goovaerts (1997)).
- Most machine learning multitask models can be placed in the context of the LMC model.

Weighted sum of Latent Functions

- In the linear model of coregionalization (LMC) outputs are expressed as linear combinations of independent random functions.
- In the LMC, each component f_d is expressed as a linear sum

$$f_d(\mathbf{x}) = \sum_{j=1}^q w_{d,j} u_j(\mathbf{x}).$$

where the latent functions are independent and have covariance functions $k_i(\mathbf{x}, \mathbf{x}')$.

• The processes $\{f_j(\mathbf{x})\}_{j=1}^q$ are independent for $q \neq j'$.

Kalman Filter Special Case

- The Kalman filter is an example of the LMC where $u_i(\mathbf{x}) \rightarrow x_i(t)$.
- I.e. we've moved form time input to a more general input space.
- In matrix notation:
 - 1. Kalman filter

 $\mathbf{F} = \mathbf{W}\mathbf{X}$

2. LMC

 $\mathbf{F} = \mathbf{W}\mathbf{U}$

where the rows of these matrices **F**, **X**, **U** each contain *q* samples from their corresponding functions at a different time (Kalman filter) or spatial location (LMC).

- If one covariance used for latent functions (like in Kalman filter).
- This is called the intrinsic coregionalization model (ICM, Goovaerts (1997)).
- The kernel matrix corresponding to a dataset **X** takes the form

- ► If outputs are noise-free, maximum likelihood is equivalent to independent fits of **B** and *k*(**x**, **x**') (Helterbrand and Cressie, 1994).
- In geostatistics this is known as autokrigeability (Wackernagel, 2003).
- In multitask learning its the cancellation of intertask transfer (Bonilla et al., 2008).

$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \mathbf{w}\mathbf{w}^{\top} \otimes k(\mathbf{X},\mathbf{X}).$$

$$\mathbf{w} = \begin{bmatrix} 1\\5 \end{bmatrix}$$
$$\mathbf{B} = \begin{bmatrix} 1 & 5\\5 & 25 \end{bmatrix}$$



$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \mathbf{w}\mathbf{w}^{\top} \otimes k(\mathbf{X},\mathbf{X}).$$





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$$\mathbf{B} = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1.5 \end{bmatrix}$$



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$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \mathbf{B}_1 \otimes k_1(\mathbf{X},\mathbf{X}) + \mathbf{B}_2 \otimes k_2(\mathbf{X},\mathbf{X})$$

$$\mathbf{B}_{1} = \begin{bmatrix} 1.4 & 0.5 \\ 0.5 & 1.2 \end{bmatrix}$$
$$\ell_{1} = 1$$
$$\mathbf{B}_{2} = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1.3 \end{bmatrix}$$
$$\ell_{2} = 0.2$$



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LMC in Machine Learning and Statistics

- Used in machine learning for GPs for multivariate regression and in statistics for computer emulation of expensive multivariate computer codes.
- Imposes the correlation of the outputs explicitly through the set of coregionalization matrices.
- Setting B = I_p assumes outputs are conditionally independent given the parameters θ. (Minka and Picard, 1997; Lawrence and Platt, 2004; Yu et al., 2005).
- More recent approaches for multiple output modeling are different versions of the linear model of coregionalization.

Semiparametric Latent Factor Model

 Coregionalization matrices are rank 1 Teh et al. (2005). rewrite equation (??) as

$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \sum_{j=1}^{q} \mathbf{w}_{:,j} \mathbf{w}_{:,j}^{\top} \otimes k_{j}(\mathbf{X},\mathbf{X}).$$

- Like the Kalman filter, but each latent function has a different covariance.
- Authors suggest using an exponentiated quadratic characteristic length-scale for each input dimension.

$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \mathbf{w}_{:,1}\mathbf{w}_{:,1}^{\top} \otimes k_1(\mathbf{X},\mathbf{X}) + \mathbf{w}_{:,2}\mathbf{w}_{:,2}^{\top} \otimes k_2(\mathbf{X},\mathbf{X})$$

$$\mathbf{w}_1 = \begin{bmatrix} 0.5\\1 \end{bmatrix}$$
$$\mathbf{w}_2 = \begin{bmatrix} 1\\0.5 \end{bmatrix}$$



$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \mathbf{w}_{:,1}\mathbf{w}_{:,1}^{\top} \otimes k_1(\mathbf{X},\mathbf{X}) + \mathbf{w}_{:,2}\mathbf{w}_{:,2}^{\top} \otimes k_2(\mathbf{X},\mathbf{X})$$





$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \mathbf{w}_{:,1}\mathbf{w}_{:,1}^{\top} \otimes k_1(\mathbf{X},\mathbf{X}) + \mathbf{w}_{:,2}\mathbf{w}_{:,2}^{\top} \otimes k_2(\mathbf{X},\mathbf{X})$$





$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \mathbf{w}_{:,1}\mathbf{w}_{:,1}^{\top} \otimes k_1(\mathbf{X},\mathbf{X}) + \mathbf{w}_{:,2}\mathbf{w}_{:,2}^{\top} \otimes k_2(\mathbf{X},\mathbf{X})$$





$$\mathbf{K}(\mathbf{X},\mathbf{X}) = \mathbf{w}_{:,1}\mathbf{w}_{:,1}^{\top} \otimes k_1(\mathbf{X},\mathbf{X}) + \mathbf{w}_{:,2}\mathbf{w}_{:,2}^{\top} \otimes k_2(\mathbf{X},\mathbf{X})$$





Gaussian processes for Multi-task, Multi-output and Multi-class

- Bonilla et al. (2008) suggest ICM for multitask learning.
- ► Use a PPCA form for **B**: similar to our Kalman filter example.
- Refer to the autokrigeability effect as the cancellation of inter-task transfer.
- Also discuss the similarities between the multi-task GP and the ICM, and its relationship to the SLFM and the LMC.

Multitask Classification

- Mostly restricted to the case where the outputs are conditionally independent given the hyperparameters φ (Minka and Picard, 1997; Williams and Barber, 1998; Lawrence and Platt, 2004; Seeger and Jordan, 2004; Yu et al., 2005; Rasmussen and Williams, 2006).
- Intrinsic coregionalization model has been used in the multiclass scenario. Skolidis and Sanguinetti (2011) use the intrinsic coregionalization model for classification, by introducing a probit noise model as the likelihood.
- Posterior distribution is no longer analytically tractable: approximate inference is required.
- A statistical model used as a surrogate for a computationally expensive computer model.
- Higdon et al. (2008) use the linear model of coregionalization to model images representing the evolution of the implosion of steel cylinders.
- In Conti and O'Hagan (2009) use the ICM to model a vegetation model: called the Sheffield Dynamic Global Vegetation Model (Woodward et al., 1998).

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