

# Fitting Covariance and Multioutput Gaussian Processes

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GPSS  
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# Outline

Parametric Models are a Bottleneck

Constructing Covariance

GP Limitations

Kalman Filter

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# Nonparametric Gaussian Processes

- ▶ We've seen how we go from parametric to non-parametric.
- ▶ The limit implies infinite dimensional  $\mathbf{w}$ .
- ▶ Gaussian processes are generally non-parametric: combine data with covariance function to get model.
- ▶ This representation *cannot* be summarized by a parameter vector of a fixed size.

# The Parametric Bottleneck

- ▶ Parametric models have a representation that does not respond to increasing training set size.
- ▶ Bayesian posterior distributions over parameters contain the information about the training data.
  - ▶ Use Bayes' rule from training data,  $p(\mathbf{w}|\mathbf{y}, \mathbf{X})$ ,
  - ▶ Make predictions on test data

$$p(y_*|\mathbf{X}_*, \mathbf{y}, \mathbf{X}) = \int p(y_*|\mathbf{w}, \mathbf{X}_*) p(\mathbf{w}|\mathbf{y}, \mathbf{X}) d\mathbf{w}.$$

- ▶  $\mathbf{w}$  becomes a bottleneck for information about the training set to pass to the test set.
- ▶ Solution: increase  $m$  so that the bottleneck is so large that it no longer presents a problem.
- ▶ How big is big enough for  $m$ ? Non-parametrics says  $m \rightarrow \infty$ .

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- ▶ These are known as degenerate covariance matrices.
- ▶ Their rank is at most  $m$ , non-parametric models have full rank covariance matrices.
- ▶ Most well known is the “linear kernel”,  $k(\mathbf{x}_i, \mathbf{x}_j) = \mathbf{x}_i^\top \mathbf{x}_j$ .

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- ▶ Parametric is a special case when conditional prediction can be summarized in a *fixed* number of parameters.
- ▶ Complexity of parametric model remains fixed regardless of the size of our training data set.
- ▶ For a non-parametric model the required number of parameters grows with the size of the training data.

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- ▶ the kernel perspective does not make a probabilistic interpretation of the covariance function.
- ▶ Algorithms can be simpler, but probabilistic interpretation is crucial for kernel parameter optimization.

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# Constructing Covariance Functions

- ▶ Sum of two covariances is also a covariance function.

$$k(\mathbf{x}, \mathbf{x}') = k_1(\mathbf{x}, \mathbf{x}') + k_2(\mathbf{x}, \mathbf{x}')$$

# Constructing Covariance Functions

- ▶ Product of two covariances is also a covariance function.

$$k(\mathbf{x}, \mathbf{x}') = k_1(\mathbf{x}, \mathbf{x}')k_2(\mathbf{x}, \mathbf{x}')$$

# Multiply by Deterministic Function

- ▶ If  $f(\mathbf{x})$  is a Gaussian process.
- ▶  $g(\mathbf{x})$  is a deterministic function.
- ▶  $h(\mathbf{x}) = f(\mathbf{x})g(\mathbf{x})$
- ▶ Then

$$k_h(\mathbf{x}, \mathbf{x}') = g(\mathbf{x})k_f(\mathbf{x}, \mathbf{x}')g(\mathbf{x}')$$

where  $k_h$  is covariance for  $h(\cdot)$  and  $k_f$  is covariance for  $f(\cdot)$ .

# Covariance Functions

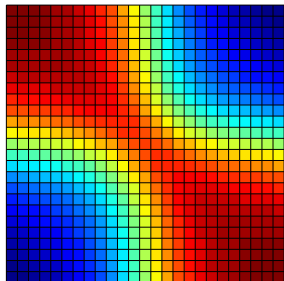
## MLP Covariance Function

$$k(\mathbf{x}, \mathbf{x}') = \alpha \sin\left(\frac{w\mathbf{x}^\top \mathbf{x}' + b}{\sqrt{w\mathbf{x}^\top \mathbf{x} + b + 1} \sqrt{w\mathbf{x}'^\top \mathbf{x}' + b + 1}}\right)$$

- Based on infinite neural network model.

$$w = 40$$

$$b = 4$$



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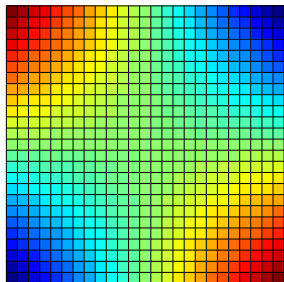
# Covariance Functions

## Linear Covariance Function

$$k(\mathbf{x}, \mathbf{x}') = \alpha \mathbf{x}^\top \mathbf{x}'$$

- Bayesian linear regression.

$$\alpha = 1$$



# Covariance Functions

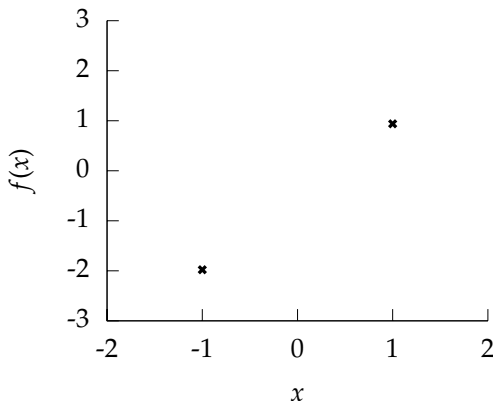
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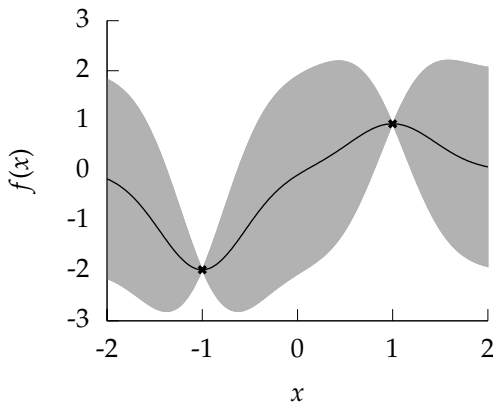
$$\alpha = 1$$

# Gaussian Process Interpolation



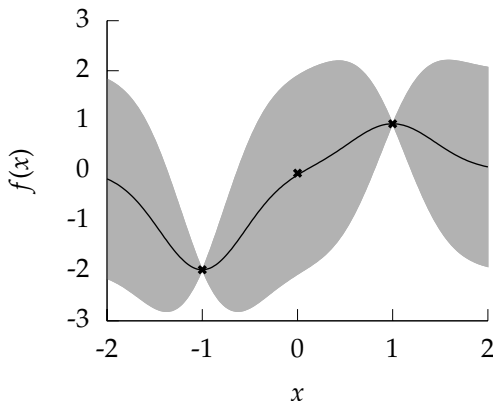
**Figure:** Real example: BACCO (see *e.g.* (Oakley and O'Hagan, 2002)).  
Interpolation through outputs from slow computer simulations (*e.g.* atmospheric carbon levels).

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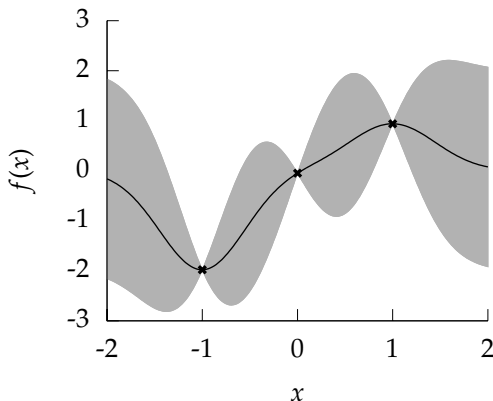
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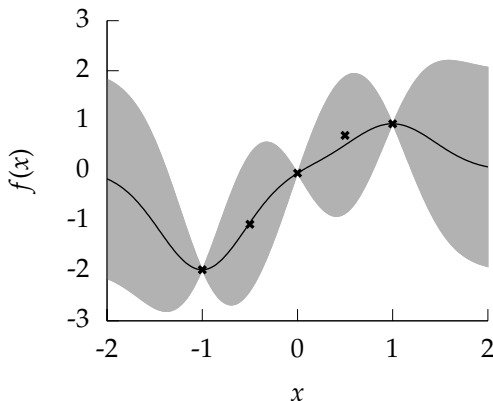
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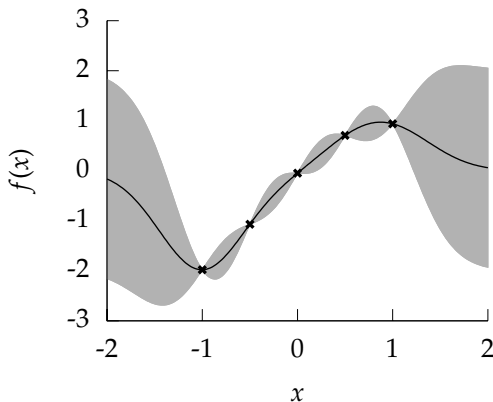
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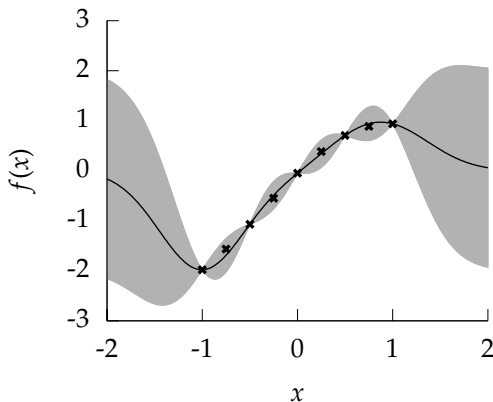
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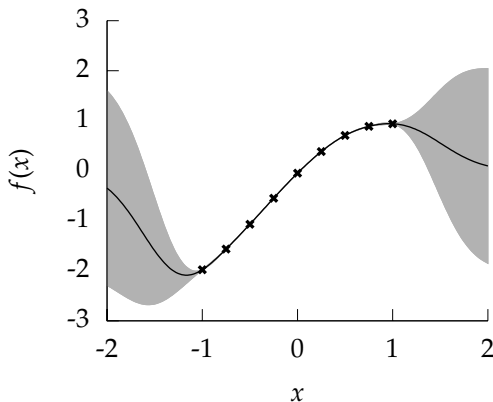


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# Gaussian Noise

- ▶ Gaussian noise model,

$$p(y_i|f_i) = \mathcal{N}(y_i|f_i, \sigma^2)$$

where  $\sigma^2$  is the variance of the noise.

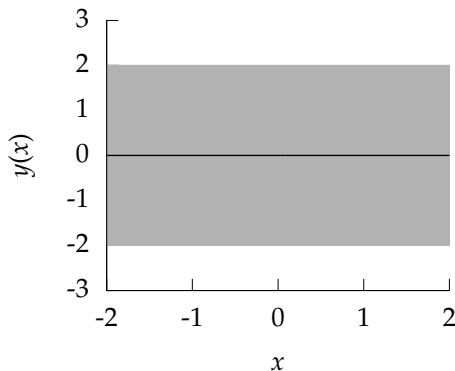
- ▶ Equivalent to a covariance function of the form

$$k(\mathbf{x}_i, \mathbf{x}_j) = \delta_{i,j} \sigma^2$$

where  $\delta_{i,j}$  is the Kronecker delta function.

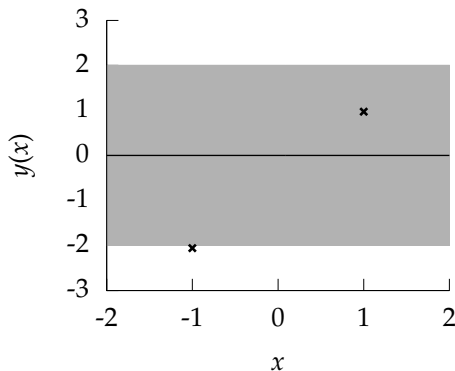
- ▶ Additive nature of Gaussians means we can simply add this term to existing covariance matrices.

# Gaussian Process Regression



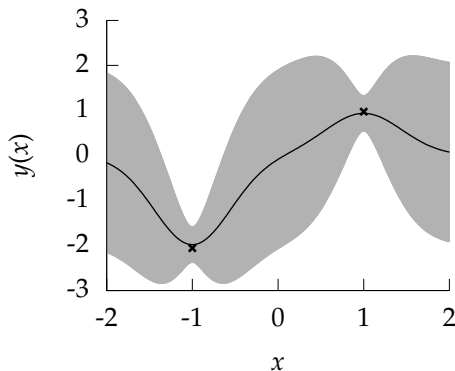
**Figure:** Examples include WiFi localization, C14 calibration curve.

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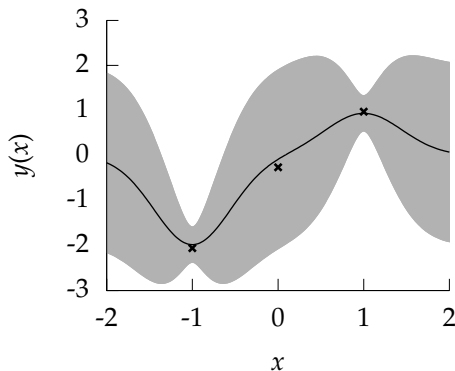
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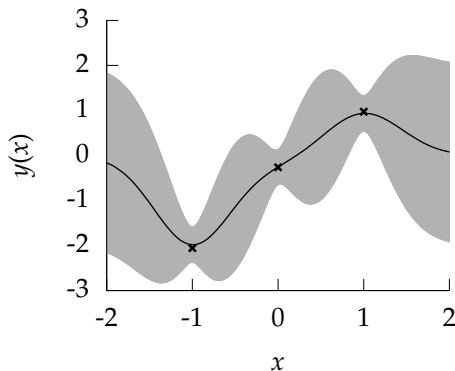
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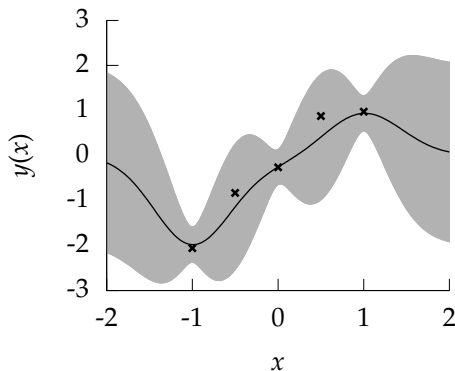
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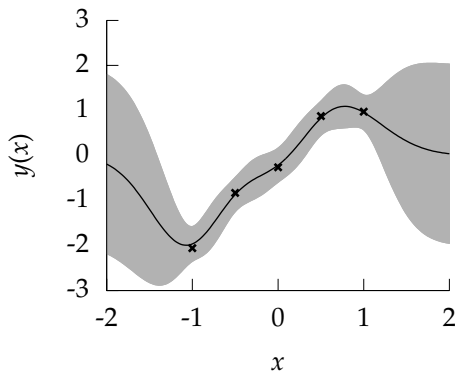


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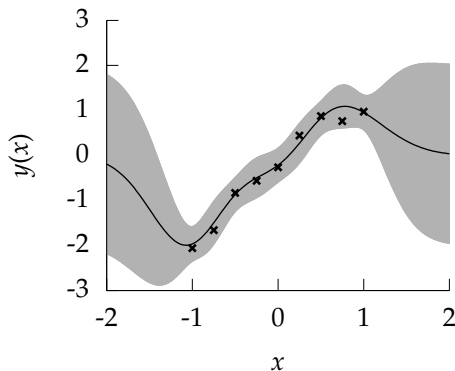
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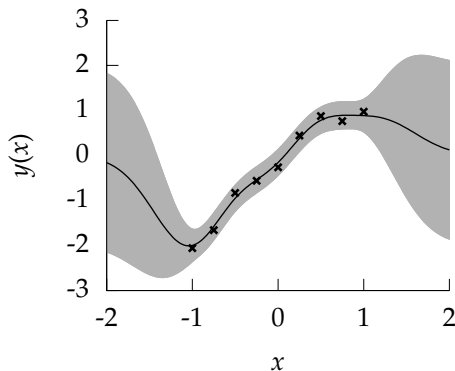
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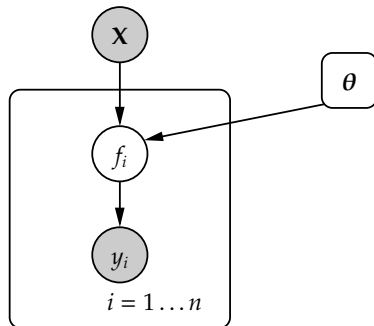


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# General Noise Models

## Graph of a GP

- ▶ Relates input variables,  $\mathbf{X}$ , to vector,  $\mathbf{y}$ , through  $\mathbf{f}$  given kernel parameters  $\theta$ .
- ▶ Plate notation indicates independence of  $y_i|f_i$ .
- ▶ In general  $p(y_i|f_i)$  is non-Gaussian.
- ▶ We approximate with Gaussian  
 $p(y_i|f_i) \approx \mathcal{N}(m_i|f_i, \beta_i^{-1})$ .



**Figure:** The Gaussian process depicted graphically.

# Gaussian Noise

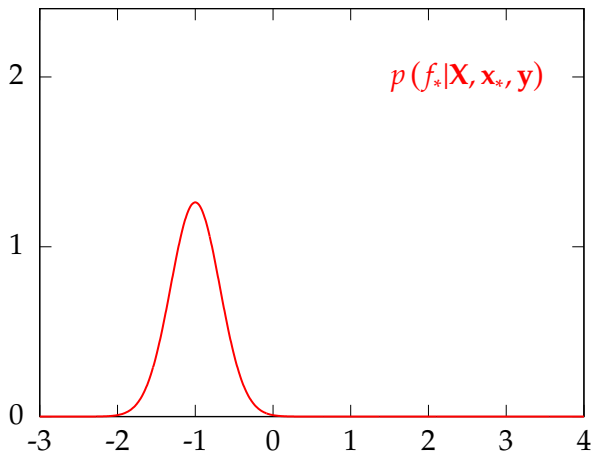


Figure: Inclusion of a data point with Gaussian noise.

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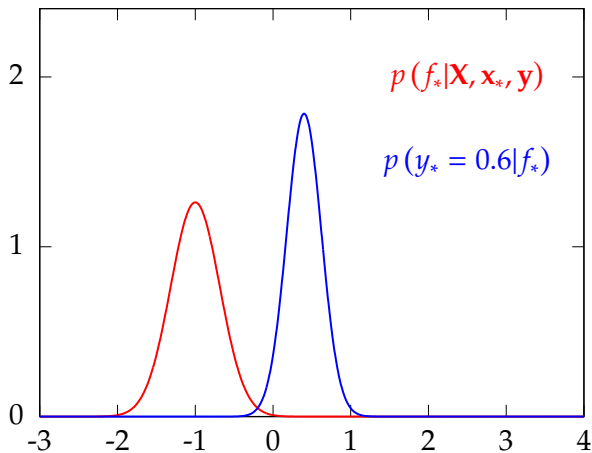


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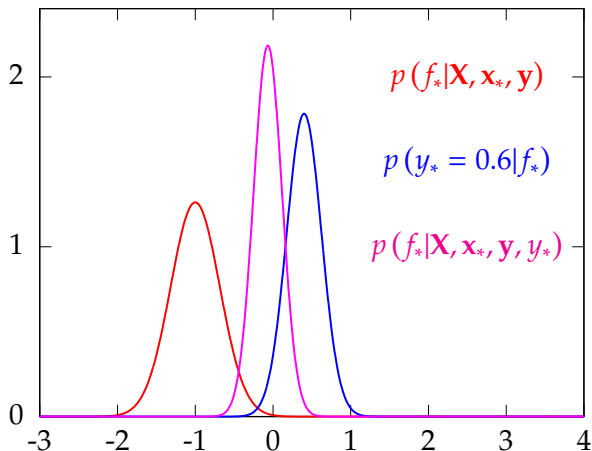


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# Expectation Propagation

## Local Moment Matching

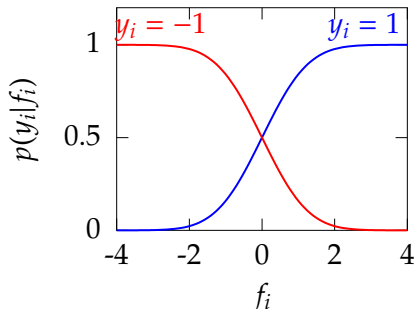
- ▶ Easiest to consider a single previously unseen data point,  $y_*, \mathbf{x}_*$ .
- ▶ Before seeing data point, prediction of  $f_*$  is a GP,  $q(f_*|\mathbf{y}, \mathbf{X})$ .
- ▶ Update prediction using Bayes' Rule,

$$p(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*) = \frac{p(y_*|f_*) p(f_*|\mathbf{y}, \mathbf{X}, \mathbf{x}_*)}{p(\mathbf{y}, y_*|\mathbf{X}, \mathbf{x}_*)}.$$

This posterior is not a Gaussian process if  $p(y_*|f_*)$  is non-Gaussian.

# Classification Noise Model

## Probit Noise Model



**Figure:** The probit model (classification). The plot shows  $p(y_i|f_i)$  for different values of  $y_i$ . For  $y_i = 1$  we have

$$p(y_i|f_i) = \Phi(f_i) = \int_{-\infty}^{f_i} \mathcal{N}(z|0, 1) dz.$$

# Expectation Propagation II

## Match Moments

- ▶ Idea behind EP — approximate with a Gaussian process at this stage by matching moments.
- ▶ This is equivalent to minimizing the following KL divergence where  $q(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*)$  is constrained to be a GP.

$$q(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*) = \operatorname{argmin}_{q(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*)} \operatorname{KL}(p(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*) \| q(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*))$$

- ▶ This is equivalent to setting

$$\langle f_* \rangle_{q(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*)} = \langle f_* \rangle_{p(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*)}$$

$$\langle f_*^2 \rangle_{q(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*)} = \langle f_*^2 \rangle_{p(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*)}$$

# Expectation Propagation III

## Equivalent Gaussian

- This is achieved by replacing  $p(y_*|f_*)$  with a Gaussian distribution

$$p(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*) = \frac{p(y_*|f_*) p(f_*|\mathbf{y}, \mathbf{X}, \mathbf{x}_*)}{p(\mathbf{y}, y_*|\mathbf{X}, \mathbf{x}_*)}$$

becomes

$$q(f_*|\mathbf{y}, y_*, \mathbf{X}, \mathbf{x}_*) = \frac{\mathcal{N}(m_*|f_*, \beta_m^{-1}) p(f_*|\mathbf{y}, \mathbf{X}, \mathbf{x}_*)}{p(\mathbf{y}, y_*|\mathbf{X}, \mathbf{x}_*)}.$$

# Classification

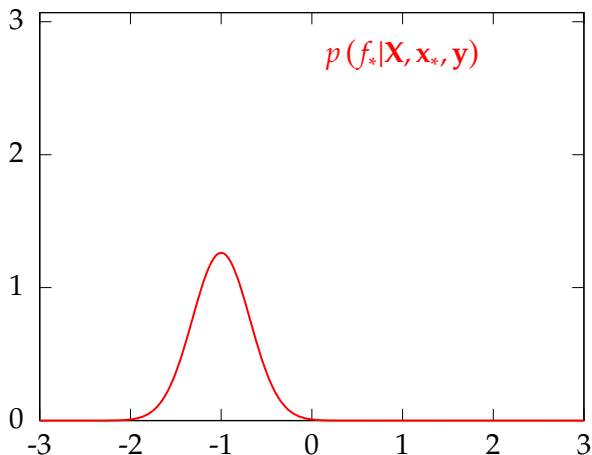


Figure: An EP style update with a classification noise model.

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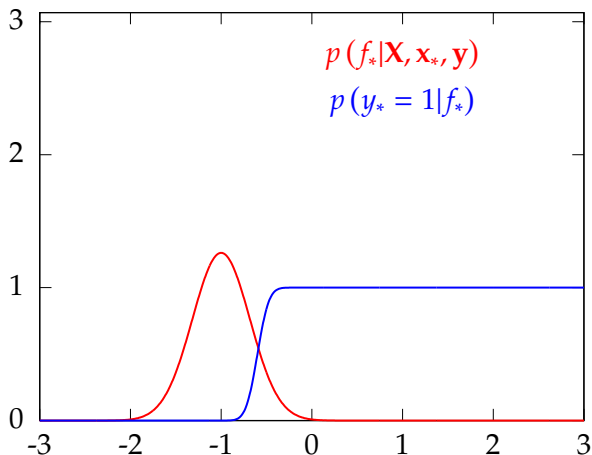


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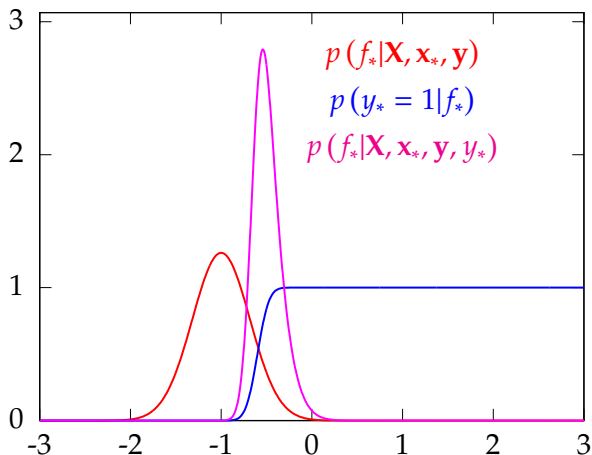


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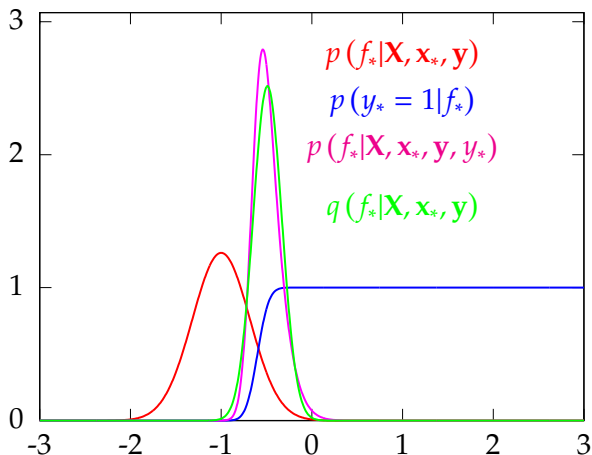
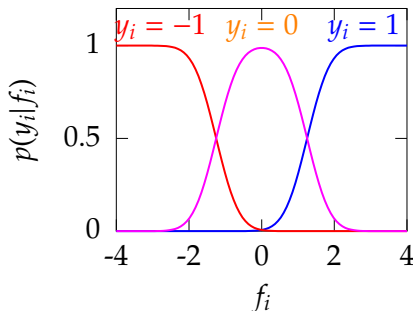


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# Ordinal Noise Model

## Ordered Categories



**Figure:** The ordered categorical noise model (ordinal regression). The plot shows  $p(y_i|f_i)$  for different values of  $y_i$ . Here we have assumed three categories.

# Laplace Approximation

- ▶ Equivalent Gaussian is found by making a local 2nd order Taylor approximation at the mode.
- ▶ Laplace was the first to suggest this<sup>1</sup>, so it's known as the *Laplace approximation*.

# Learning Covariance Parameters

Can we determine covariance parameters from the data?

$$\mathcal{N}(\mathbf{y}|\mathbf{0}, \mathbf{K}) = \frac{1}{(2\pi)^{\frac{n}{2}} |\mathbf{K}|^{\frac{1}{2}}} \exp\left(-\frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}\right)$$

The parameters are *inside* the covariance function (matrix).

$$k_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j; \theta)$$

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Can we determine covariance parameters from the data?

$$\log \mathcal{N}(\mathbf{y}|\mathbf{0}, \mathbf{K}) = -\frac{1}{2} \log |\mathbf{K}| - \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2} - \frac{n}{2} \log 2\pi$$

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$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

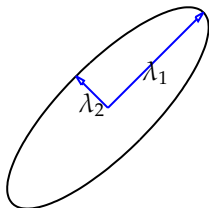
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# Eigendecomposition of Covariance

A useful decomposition for understanding the objective function.

$$\mathbf{K} = \mathbf{R}\mathbf{\Lambda}^2\mathbf{R}^\top$$

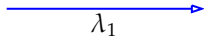


Diagonal of  $\mathbf{\Lambda}$  represents distance along axes.

$\mathbf{R}$  gives a rotation of these axes.

## Capacity control: $\log |\mathbf{K}|$

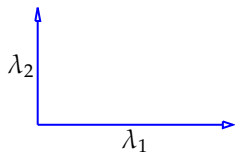
$$\mathbf{\Lambda} = \begin{bmatrix} \boxed{\lambda_1 & 0} \\ 0 & \lambda_2 \end{bmatrix}$$





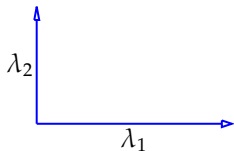
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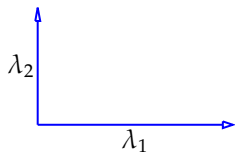
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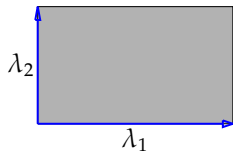
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$$|\mathbf{\Lambda}| = \lambda_1 \lambda_2$$

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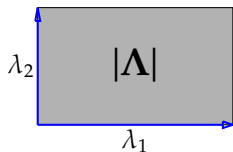
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## Capacity control: $\log |\mathbf{K}|$

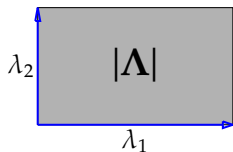
$$\mathbf{\Lambda} = \begin{bmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{bmatrix}$$



$$|\mathbf{\Lambda}| = \lambda_1 \lambda_2$$

## Capacity control: $\log |\mathbf{K}|$

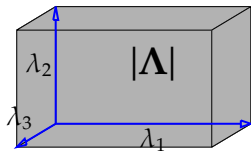
$$\mathbf{\Lambda} = \begin{bmatrix} \lambda_1 & 0 & 0 \\ 0 & \lambda_2 & 0 \\ 0 & 0 & \lambda_3 \end{bmatrix}$$



$$|\mathbf{\Lambda}| = \lambda_1 \lambda_2$$

## Capacity control: $\log |\mathbf{K}|$

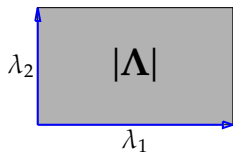
$$\mathbf{\Lambda} = \begin{bmatrix} \lambda_1 & 0 & 0 \\ 0 & \lambda_2 & 0 \\ 0 & 0 & \lambda_3 \end{bmatrix}$$



$$|\mathbf{\Lambda}| = \lambda_1 \lambda_2 \lambda_3$$

## Capacity control: $\log |\mathbf{K}|$

$$\mathbf{\Lambda} = \begin{bmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{bmatrix}$$

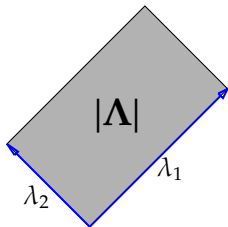


$$|\mathbf{\Lambda}| = \lambda_1 \lambda_2$$



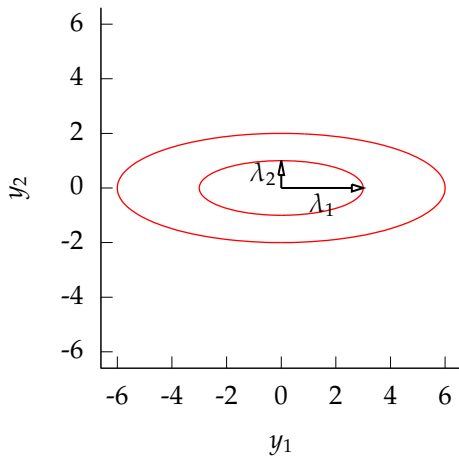
## Capacity control: $\log |\mathbf{K}|$

$$\mathbf{R}\mathbf{\Lambda} = \begin{bmatrix} w_{1,1} & w_{1,2} \\ w_{2,1} & w_{2,2} \end{bmatrix}$$

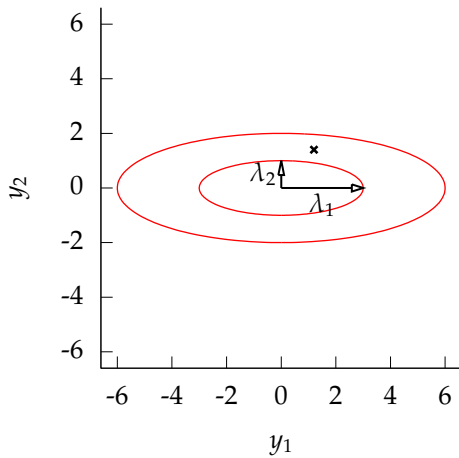


$$|\mathbf{R}\mathbf{\Lambda}| = \lambda_1 \lambda_2$$

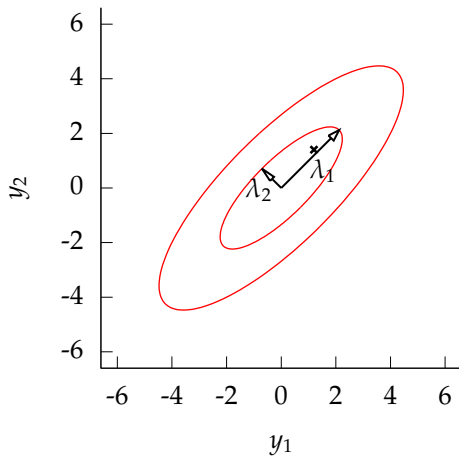
Data Fit:  $\frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$



Data Fit:  $\frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$

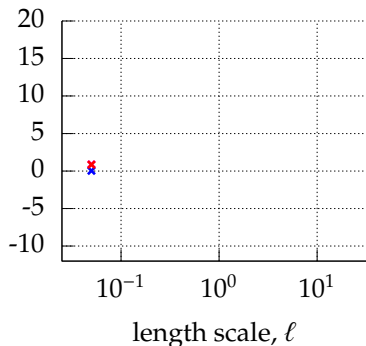
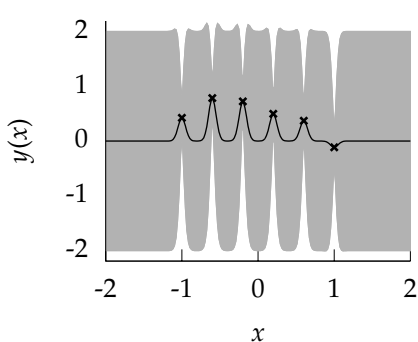


Data Fit:  $\frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$



# Learning Covariance Parameters

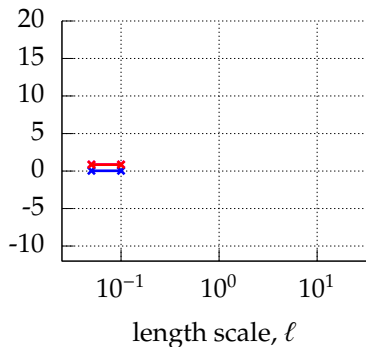
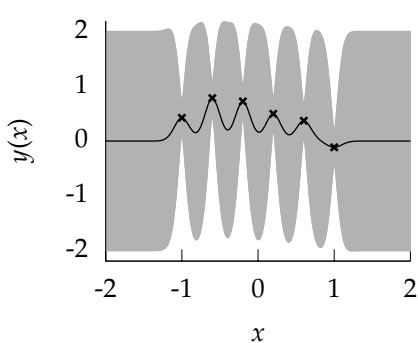
Can we determine length scales and noise levels from the data?



$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

# Learning Covariance Parameters

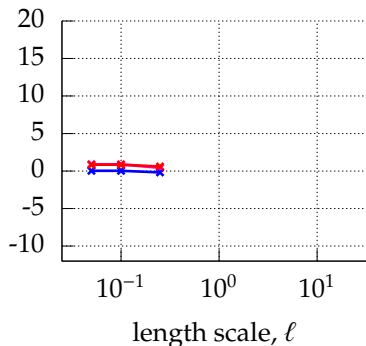
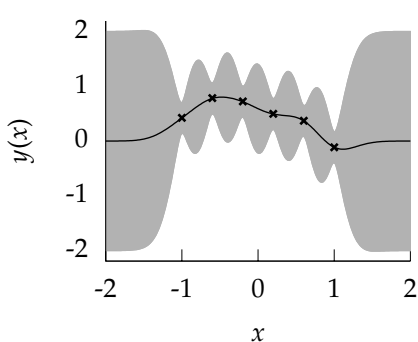
Can we determine length scales and noise levels from the data?



$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

# Learning Covariance Parameters

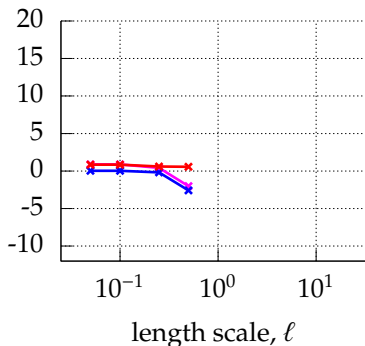
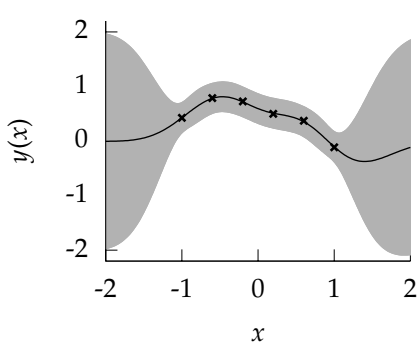
Can we determine length scales and noise levels from the data?



$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

# Learning Covariance Parameters

Can we determine length scales and noise levels from the data?

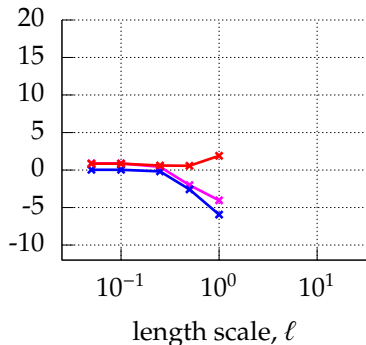
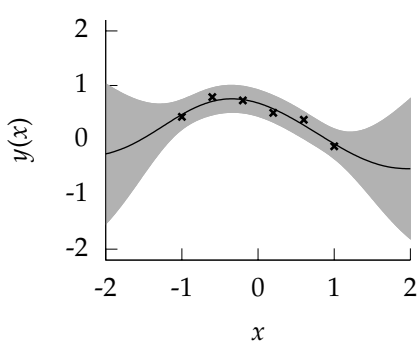


$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$



# Learning Covariance Parameters

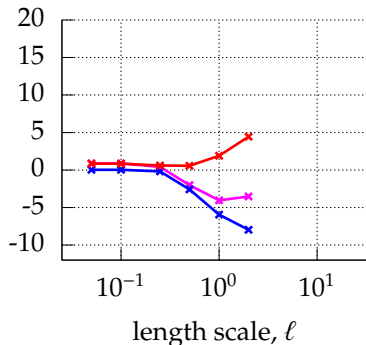
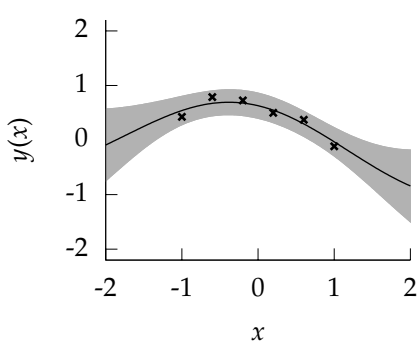
Can we determine length scales and noise levels from the data?



$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

# Learning Covariance Parameters

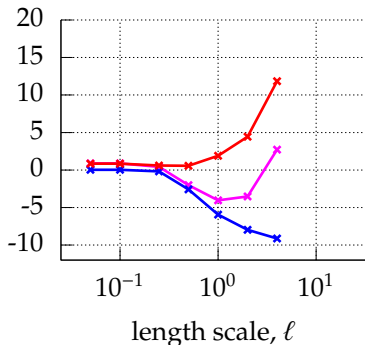
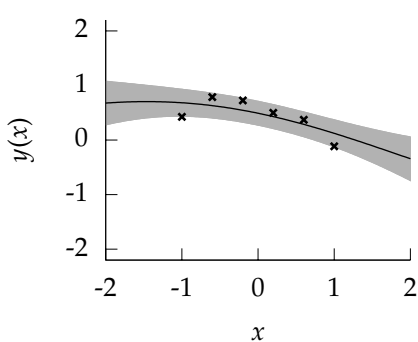
Can we determine length scales and noise levels from the data?



$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

# Learning Covariance Parameters

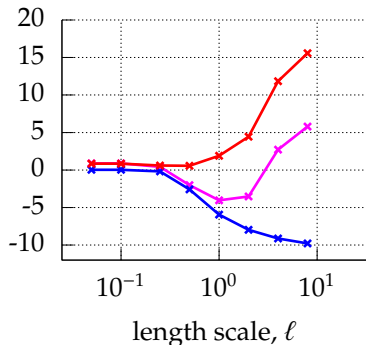
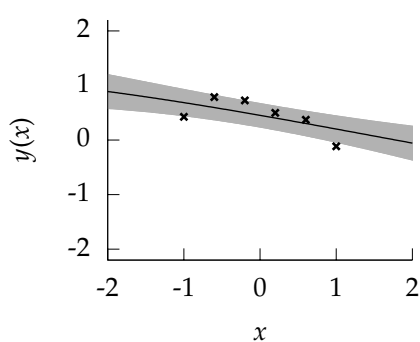
Can we determine length scales and noise levels from the data?



$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

# Learning Covariance Parameters

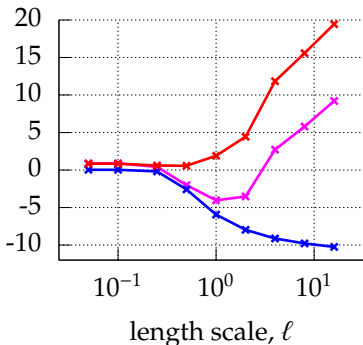
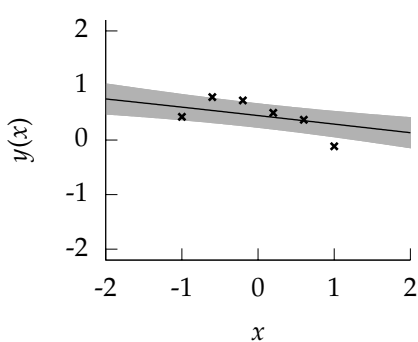
Can we determine length scales and noise levels from the data?



$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

# Learning Covariance Parameters

Can we determine length scales and noise levels from the data?



$$E(\theta) = \frac{1}{2} \log |\mathbf{K}| + \frac{\mathbf{y}^\top \mathbf{K}^{-1} \mathbf{y}}{2}$$

# Gene Expression Example

- ▶ Given given expression levels in the form of a time series from Della Gatta et al. (2008).
- ▶ Want to detect if a gene is expressed or not, fit a GP to each gene (Kalaitzis and Lawrence, 2011).

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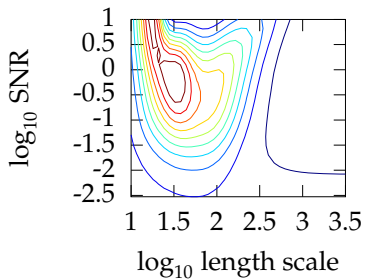
# A Simple Approach to Ranking Differentially Expressed Gene Expression Time Courses through Gaussian Process Regression

Alfredo A Kalaitzis\* and Neil D Lawrence\*

## Abstract

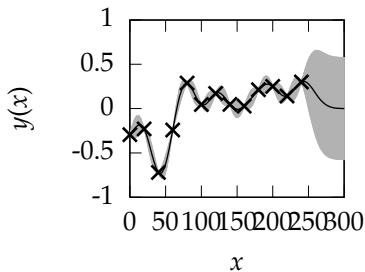
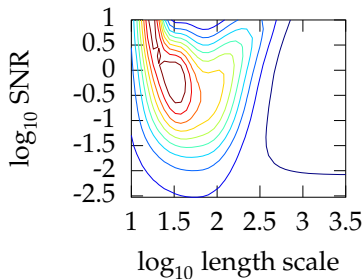
**Background:** The analysis of gene expression from time series underpins many biological studies. Two basic forms of analysis recur for data of this type: removing inactive (quiet) genes from the study and determining which genes are differentially expressed. Often these analysis stages are applied disregarding the fact that the data is drawn from a time series. In this paper we propose a simple model for accounting for the underlying temporal nature of the data based on a Gaussian process.

**Results:** We review Gaussian process (GP) regression for estimating the continuous trajectories underlying in gene expression time-series. We present a simple approach which can be used to filter quiet genes, or for the case of time series in the form of expression ratios, quantify differential expression. We assess via ROC curves the rankings produced by our regression framework and compare them to a recently proposed hierarchical Bayesian model for the analysis of gene expression time-series (BATS). We compare on both simulated and experimental data showing that the proposed approach considerably outperforms the current state of the art.

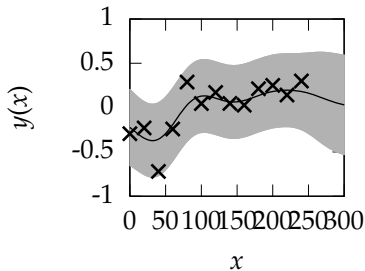
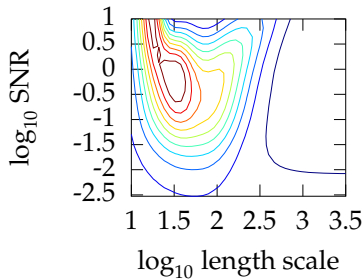


Contour plot of Gaussian process likelihood.

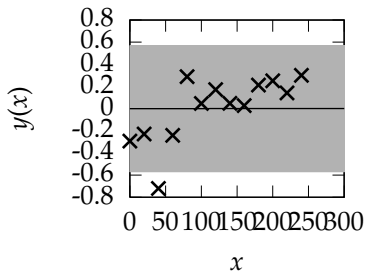
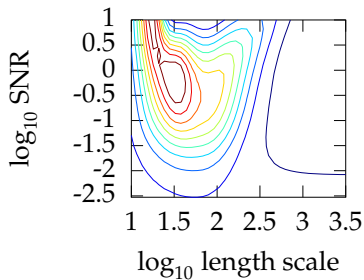




Optima: length scale of 1.2221 and  $\log_{10}$  SNR of 1.9654  
 log likelihood is -0.22317.



Optima: length scale of 1.5162 and  $\log_{10}$  SNR of 0.21306  
 log likelihood is -0.23604.



Optima: length scale of 2.9886 and  $\log_{10}$  SNR of -4.506  
log likelihood is -2.1056.

# Outline

Parametric Models are a Bottleneck

Constructing Covariance

**GP Limitations**

Kalman Filter

# Limitations of Gaussian Processes

- ▶ Inference is  $O(n^3)$  due to matrix inverse (in practice use Cholesky).
- ▶ Gaussian processes don't deal well with discontinuities (financial crises, phosphorylation, collisions, edges in images).
- ▶ Widely used exponentiated quadratic covariance (RBF) can be too smooth in practice (but there are many alternatives!!).

# Outline

Parametric Models are a Bottleneck

Constructing Covariance

GP Limitations

**Kalman Filter**

# Simple Markov Chain

- ▶ Assume 1-d latent state, a vector over time,  $\mathbf{x} = [x_1 \dots x_T]$ .
- ▶ Markov property,

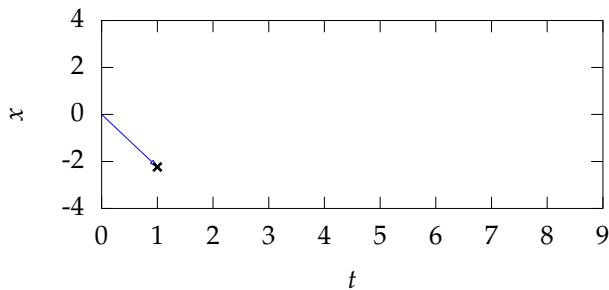
$$\begin{aligned}x_i &= x_{i-1} + \epsilon_i, \\ \epsilon_i &\sim \mathcal{N}(0, \alpha) \\ \implies x_i &\sim \mathcal{N}(x_{i-1}, \alpha)\end{aligned}$$

- ▶ Initial state,

$$x_0 \sim \mathcal{N}(0, \alpha_0)$$

- ▶ If  $x_0 \sim \mathcal{N}(0, \alpha)$  we have a Markov chain for the latent states.
- ▶ Markov chain it is specified by an initial distribution (Gaussian) and a transition distribution (Gaussian).

# Gauss Markov Chain



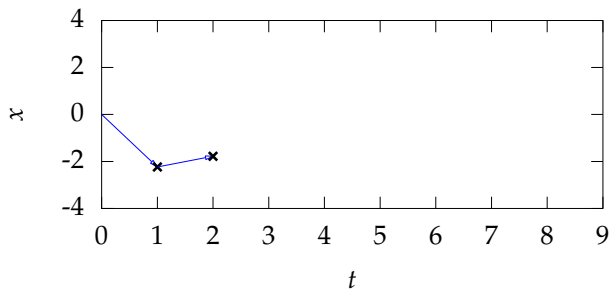
$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_0 = 0.000, \quad \epsilon_1 = -2.24$$

$$x_1 = 0.000 - 2.24 = -2.24$$



# Gauss Markov Chain

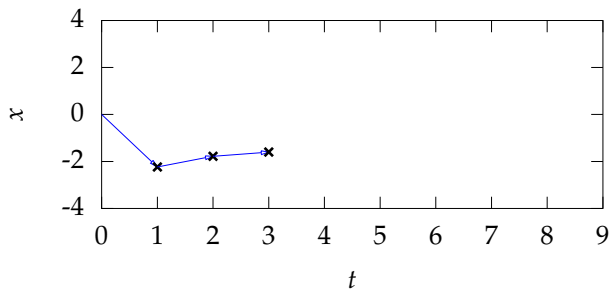


$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_1 = -2.24, \quad \epsilon_2 = 0.457$$

$$x_2 = -2.24 + 0.457 = -1.78$$

# Gauss Markov Chain

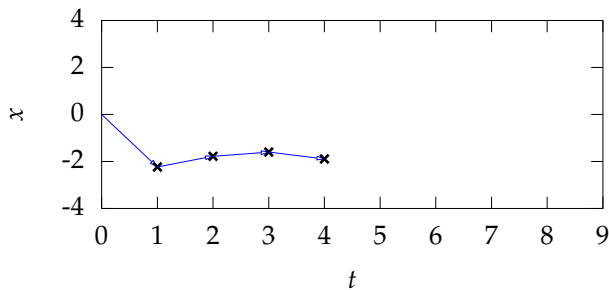


$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_2 = -1.78, \quad \epsilon_3 = 0.178$$

$$x_3 = -1.78 + 0.178 = -1.6$$

# Gauss Markov Chain

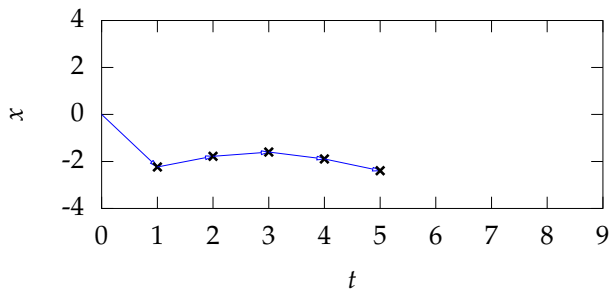


$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_3 = -1.6, \quad \epsilon_4 = -0.292$$

$$x_4 = -1.6 - 0.292 = -1.89$$

# Gauss Markov Chain

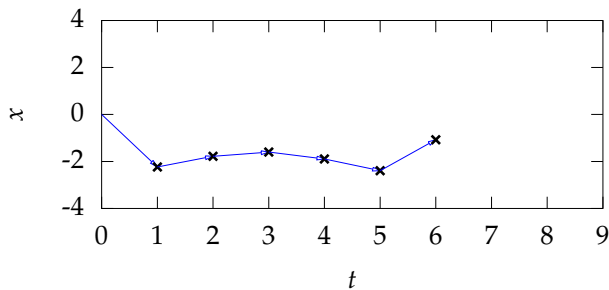


$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_4 = -1.89, \quad \epsilon_5 = -0.501$$

$$x_5 = -1.89 - 0.501 = -2.39$$

# Gauss Markov Chain

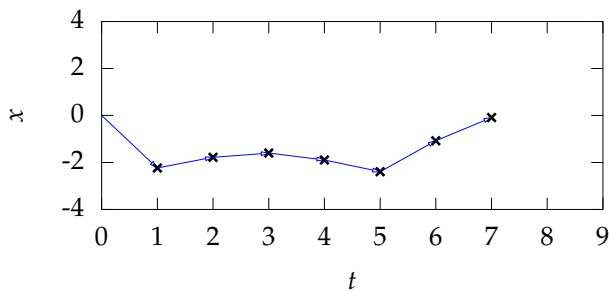


$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_5 = -2.39, \quad \epsilon_6 = 1.32$$

$$x_6 = -2.39 + 1.32 = -1.08$$

# Gauss Markov Chain

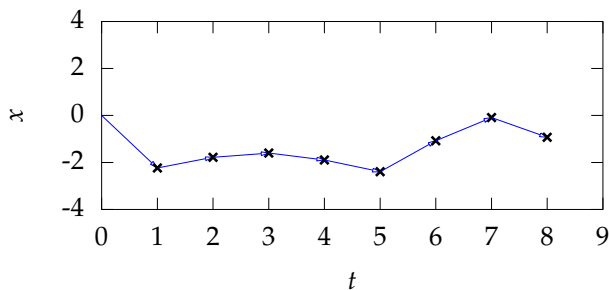


$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_6 = -1.08, \quad \epsilon_7 = 0.989$$

$$x_7 = -1.08 + 0.989 = -0.0881$$

# Gauss Markov Chain

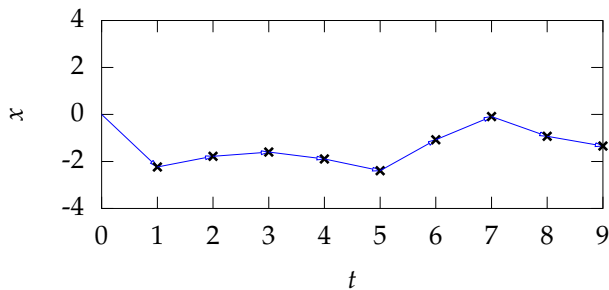


$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_7 = -0.0881, \quad \epsilon_8 = -0.842$$

$$x_8 = -0.0881 - 0.842 = -0.93$$

# Gauss Markov Chain



$$x_0 = 0, \quad \epsilon_i \sim \mathcal{N}(0, 1)$$

$$x_8 = -0.93, \quad \epsilon_9 = -0.41$$

$$x_9 = -0.93 - 0.410 = -1.34$$



# Multivariate Gaussian Properties: Reminder

If

$$\mathbf{z} \sim \mathcal{N}(\boldsymbol{\mu}, \mathbf{C})$$

and

$$\mathbf{x} = \mathbf{W}\mathbf{z} + \mathbf{b}$$

then

$$\mathbf{x} \sim \mathcal{N}(\mathbf{W}\boldsymbol{\mu} + \mathbf{b}, \mathbf{W}\mathbf{C}\mathbf{W}^\top)$$

# Multivariate Gaussian Properties: Reminder

**Simplified:** If

$$\mathbf{z} \sim \mathcal{N}(0, \sigma^2 \mathbf{I})$$

and

$$\mathbf{x} = \mathbf{W}\mathbf{z}$$

then

$$\mathbf{x} \sim \mathcal{N}(0, \sigma^2 \mathbf{W}\mathbf{W}^\top)$$

# Matrix Representation of Latent Variables

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \\ x_5 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 \\ 1 & 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 1 & 0 \\ 1 & 1 & 1 & 1 & 1 \end{bmatrix} \times \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \end{bmatrix}$$

$$x_1 = \epsilon_1$$

# Matrix Representation of Latent Variables

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \\ x_5 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 \\ 1 & 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 1 & 0 \\ 1 & 1 & 1 & 1 & 1 \end{bmatrix} \times \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \end{bmatrix}$$

$$x_2 = \epsilon_1 + \epsilon_2$$

# Matrix Representation of Latent Variables

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \\ x_5 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 \\ 1 & 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 1 & 0 \\ 1 & 1 & 1 & 1 & 1 \end{bmatrix} \times \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \end{bmatrix}$$

$$x_3 = \epsilon_1 + \epsilon_2 + \epsilon_3$$

# Matrix Representation of Latent Variables

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \\ x_5 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 \\ 1 & 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 1 & 0 \\ 1 & 1 & 1 & 1 & 1 \end{bmatrix} \times \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \end{bmatrix}$$

$$x_4 = \epsilon_1 + \epsilon_2 + \epsilon_3 + \epsilon_4$$

# Matrix Representation of Latent Variables

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \\ x_5 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 \\ 1 & 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 1 & 0 \\ 1 & 1 & 1 & 1 & 1 \end{bmatrix} \times \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \end{bmatrix}$$

$$x_5 = \epsilon_1 + \epsilon_2 + \epsilon_3 + \epsilon_4 + \epsilon_5$$

# Matrix Representation of Latent Variables

$$\mathbf{x} = \mathbf{L}_1 \times \boldsymbol{\epsilon}$$



# Multivariate Process

- ▶ Since  $\mathbf{x}$  is linearly related to  $\epsilon$  we know  $\mathbf{x}$  is a Gaussian process.
- ▶ Trick: we only need to compute the mean and covariance of  $\mathbf{x}$  to determine that Gaussian.

$$\mathbf{x} = \mathbf{L}_1 \boldsymbol{\epsilon}$$

$$\langle \mathbf{x} \rangle = \langle \mathbf{L}_1 \boldsymbol{\epsilon} \rangle$$

$$\langle \mathbf{x} \rangle = \mathbf{L}_1 \langle \epsilon \rangle$$

$$\langle \mathbf{x} \rangle = \mathbf{L}_1 \langle \boldsymbol{\epsilon} \rangle$$

$$\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{I})$$

$$\langle \mathbf{x} \rangle = \mathbf{L}_1 \mathbf{0}$$

## Latent Process Mean

$$\langle \mathbf{x} \rangle = \mathbf{0}$$

$$\mathbf{x}\mathbf{x}^\top = \mathbf{L}_1\boldsymbol{\epsilon}\boldsymbol{\epsilon}^\top\mathbf{L}_1^\top$$

$$\mathbf{x}^\top = \boldsymbol{\epsilon}^\top\mathbf{L}^\top$$



## Latent Process Covariance

$$\langle \mathbf{x} \mathbf{x}^\top \rangle = \langle \mathbf{L}_1 \boldsymbol{\epsilon} \boldsymbol{\epsilon}^\top \mathbf{L}_1^\top \rangle$$

## Latent Process Covariance

$$\langle \mathbf{x} \mathbf{x}^\top \rangle = \mathbf{L}_1 \langle \boldsymbol{\epsilon} \boldsymbol{\epsilon}^\top \rangle \mathbf{L}_1^\top$$

## Latent Process Covariance

$$\langle \mathbf{x} \mathbf{x}^\top \rangle = \mathbf{L}_1 \langle \boldsymbol{\epsilon} \boldsymbol{\epsilon}^\top \rangle \mathbf{L}_1^\top$$

$$\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{I})$$

## Latent Process Covariance

$$\langle \mathbf{x} \mathbf{x}^\top \rangle = \alpha \mathbf{L}_1 \mathbf{L}_1^\top$$

$$\mathbf{x} = \mathbf{L}_1 \boldsymbol{\epsilon}$$

$$\mathbf{x} = \mathbf{L}_1 \boldsymbol{\epsilon}$$

$$\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{I})$$

$$\mathbf{x} = \mathbf{L}_1 \boldsymbol{\epsilon}$$

$$\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{I})$$

$$\implies$$

$$\mathbf{x} = \mathbf{L}_1 \boldsymbol{\epsilon}$$

$$\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{I})$$

$$\implies$$

$$\mathbf{x} \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{L}_1 \mathbf{L}_1^\top)$$



## Covariance for Latent Process II

- ▶ Make the variance dependent on time interval.
- ▶ Assume variance grows *linearly* with time.
- ▶ Justification: sum of two Gaussian distributed random variables is distributed as Gaussian with sum of variances.
- ▶ If variable's movement is additive over time (as described) variance scales linearly with time.

## Covariance for Latent Process II

- Given

$$\epsilon \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{I}) \implies \epsilon \sim \mathcal{N}(\mathbf{0}, \alpha \mathbf{L}_1 \mathbf{L}_1^\top).$$

Then

$$\epsilon \sim \mathcal{N}(\mathbf{0}, \Delta t \alpha \mathbf{I}) \implies \epsilon \sim \mathcal{N}(\mathbf{0}, \Delta t \alpha \mathbf{L}_1 \mathbf{L}_1^\top).$$

where  $\Delta t$  is the time interval between observations.

## Covariance for Latent Process II

$$\boldsymbol{\epsilon} \sim \mathcal{N}(0, \alpha \Delta t \mathbf{I}), \quad \mathbf{x} \sim \mathcal{N}(0, \alpha \Delta t \mathbf{L}_1 \mathbf{L}_1^\top)$$

## Covariance for Latent Process II

$$\boldsymbol{\epsilon} \sim \mathcal{N}(0, \alpha \Delta t \mathbf{I}), \quad \mathbf{x} \sim \mathcal{N}(0, \alpha \Delta t \mathbf{L}_1 \mathbf{L}_1^\top)$$

$$\mathbf{K} = \alpha \Delta t \mathbf{L}_1 \mathbf{L}_1^\top$$

## Covariance for Latent Process II

$$\boldsymbol{\epsilon} \sim \mathcal{N}(0, \alpha \Delta t \mathbf{I}), \quad \mathbf{x} \sim \mathcal{N}(0, \alpha \Delta t \mathbf{L}_1 \mathbf{L}_1^\top)$$

$$\mathbf{K} = \alpha \Delta t \mathbf{L}_1 \mathbf{L}_1^\top$$

$$k_{i,j} = \alpha \Delta t \mathbf{l}_{:,i}^\top \mathbf{l}_{:,j}$$

where  $\mathbf{l}_{:,k}$  is a vector from the  $k$ th row of  $\mathbf{L}_1$ : the first  $k$  elements are one, the next  $T - k$  are zero.

## Covariance for Latent Process II

$$\boldsymbol{\epsilon} \sim \mathcal{N}(0, \alpha \Delta t \mathbf{I}), \quad \mathbf{x} \sim \mathcal{N}(0, \alpha \Delta t \mathbf{L}_1 \mathbf{L}_1^\top)$$

$$\mathbf{K} = \alpha \Delta t \mathbf{L}_1 \mathbf{L}_1^\top$$

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$$k_{i,j} = \alpha \Delta t \min(i, j)$$

define  $\Delta t i = t_i$  so

$$k_{i,j} = \alpha \min(t_i, t_j) = k(t_i, t_j)$$

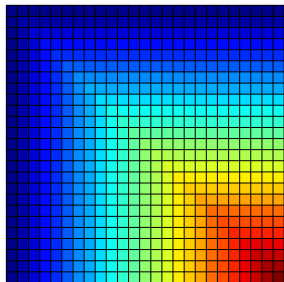
# Covariance Functions

Where did this covariance matrix come from?

## Markov Process

$$k(t, t') = \alpha \min(t, t')$$

- Covariance matrix is built using the *inputs* to the function  $t$ .



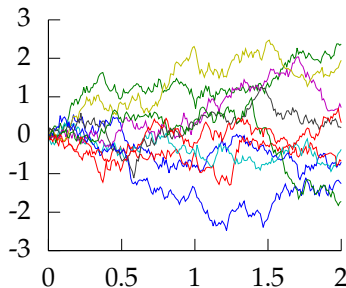
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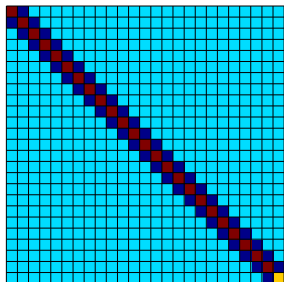
# Covariance Functions

Where did this covariance matrix come from?

## Markov Process

### Visualization of inverse covariance (precision).

- ▶ Precision matrix is sparse: only neighbours in matrix are non-zero.
- ▶ This reflects *conditional* independencies in data.
- ▶ In this case *Markov* structure.



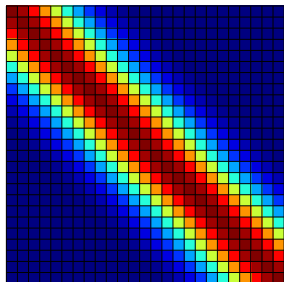
# Covariance Functions

Where did this covariance matrix come from?

## Exponentiated Quadratic Kernel Function (RBF, Squared Exponential, Gaussian)

$$k(\mathbf{x}, \mathbf{x}') = \alpha \exp\left(-\frac{\|\mathbf{x} - \mathbf{x}'\|_2^2}{2\ell^2}\right)$$

- ▶ Covariance matrix is built using the *inputs* to the function  $\mathbf{x}$ .
- ▶ For the example above it was based on Euclidean distance.
- ▶ The covariance function is also known as a kernel.



# Covariance Functions

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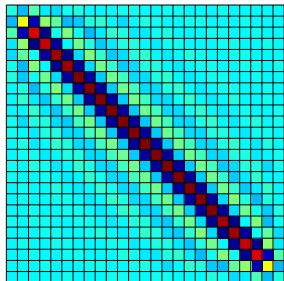
# Covariance Functions

Where did this covariance matrix come from?

## Exponentiated Quadratic

**Visualization of inverse covariance (precision).**

- ▶ Precision matrix is not sparse.
- ▶ Each point is dependent on all the others.
- ▶ In this case non-Markovian.



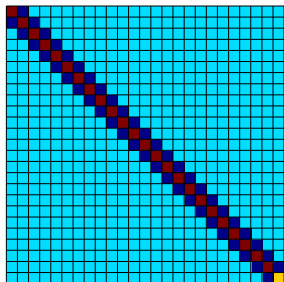
# Covariance Functions

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## Markov Process

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- ▶ In this case *Markov* structure.



# Simple Kalman Filter I

- ▶ We have state vector  $\mathbf{X} = [\mathbf{x}_1 \dots \mathbf{x}_q] \in \mathbb{R}^{T \times q}$  and if each state evolves independently we have

$$p(\mathbf{X}) = \prod_{i=1}^q p(\mathbf{x}_{:,i})$$
$$p(\mathbf{x}_{:,i}) = \mathcal{N}(\mathbf{x}_{:,i} | \mathbf{0}, \mathbf{K}).$$

- ▶ We want to obtain outputs through:

$$\mathbf{y}_{i,:} = \mathbf{W}\mathbf{x}_{i,:}$$

# Stacking and Kronecker Products I

- Represent with a 'stacked' system:

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x} | \mathbf{0}, \mathbf{I} \otimes \mathbf{K})$$

where the stacking is placing each column of  $\mathbf{X}$  one on top of another as

$$\mathbf{x} = \begin{bmatrix} \mathbf{x}_{:,1} \\ \mathbf{x}_{:,2} \\ \vdots \\ \mathbf{x}_{:,q} \end{bmatrix}$$

# Kronecker Product

$$\begin{bmatrix} a & b \\ c & d \end{bmatrix} \otimes \mathbf{K} = \begin{bmatrix} a\mathbf{K} & b\mathbf{K} \\ c\mathbf{K} & d\mathbf{K} \end{bmatrix}$$



# Kronecker Product

$$\begin{bmatrix} \text{dark gray} & \text{medium gray} \\ \text{medium gray} & \text{white} \end{bmatrix} \otimes \begin{bmatrix} \text{red} & \text{green} \\ \text{green} & \text{blue} \end{bmatrix} = \begin{bmatrix} \text{dark red} & \text{dark green} & \text{red} & \text{green} \\ \text{dark green} & \text{dark blue} & \text{green} & \text{dark blue} \\ \text{red} & \text{green} & \text{red} & \text{green} \\ \text{green} & \text{dark blue} & \text{green} & \text{blue} \end{bmatrix}$$

# Stacking and Kronecker Products I

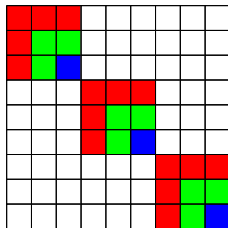
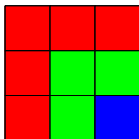
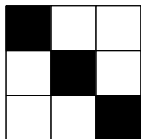
- Represent with a 'stacked' system:

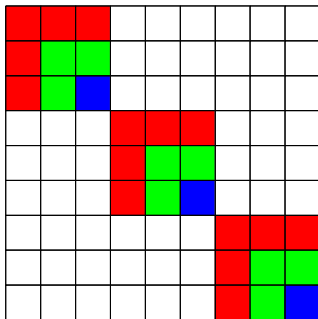
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where the stacking is placing each column of  $\mathbf{X}$  one on top of another as

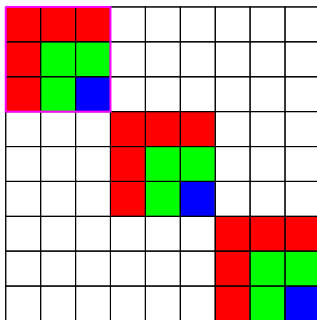
$$\mathbf{x} = \begin{bmatrix} \mathbf{x}_{:,1} \\ \mathbf{x}_{:,2} \\ \vdots \\ \mathbf{x}_{:,q} \end{bmatrix}$$

# Column Stacking

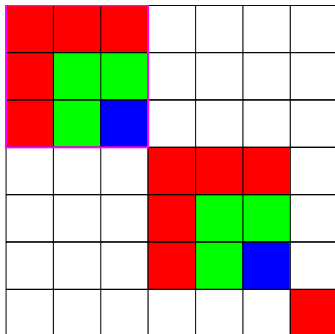




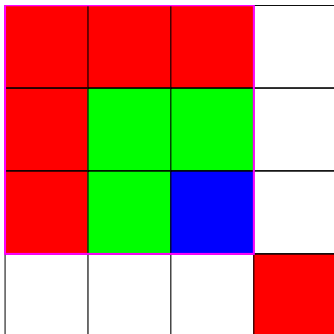
For this stacking the marginal distribution over *time* is given by the block diagonals.



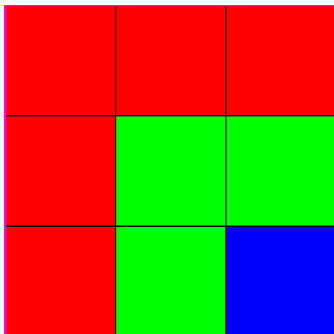
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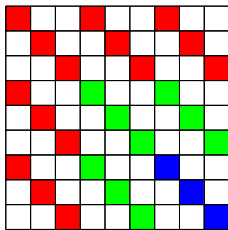
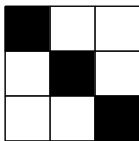
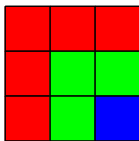
## Two Ways of Stacking

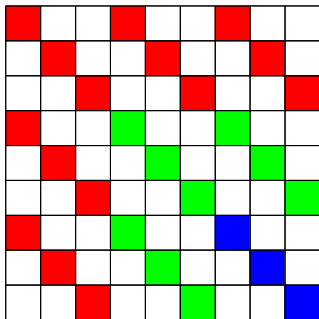
Can also stack each row of  $\mathbf{X}$  to form column vector:

$$\mathbf{x} = \begin{bmatrix} \mathbf{x}_{1,:} \\ \mathbf{x}_{2,:} \\ \vdots \\ \mathbf{x}_{T,:} \end{bmatrix}$$

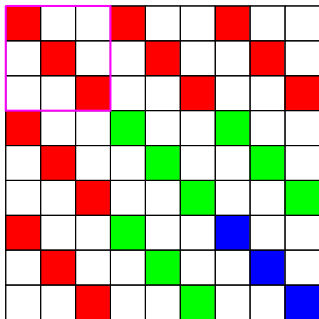
$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x} | \mathbf{0}, \mathbf{K} \otimes \mathbf{I})$$

# Row Stacking

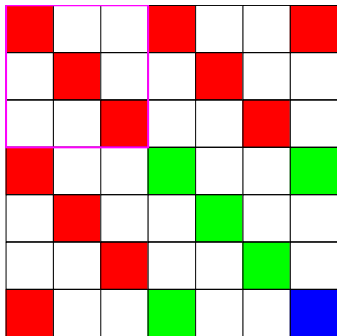




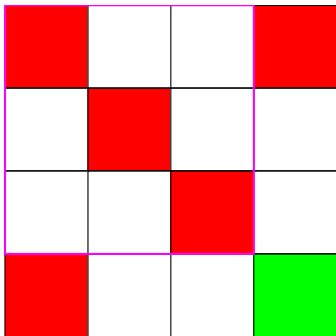
For this stacking the marginal distribution over the latent *dimensions* is given by the block diagonals.



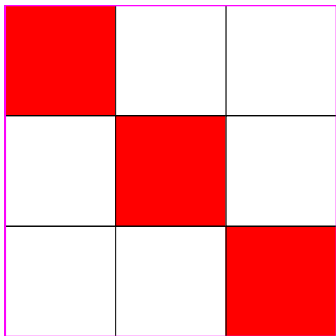
For this stacking the marginal distribution over the latent *dimensions* is given by the block diagonals.



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# Observed Process

The observations are related to the latent points by a linear mapping matrix,

$$\mathbf{y}_{i,:} = \mathbf{W}\mathbf{x}_{i,:} + \boldsymbol{\epsilon}_{i,:}$$
$$\boldsymbol{\epsilon} \sim \mathcal{N}(0, \sigma^2 \mathbf{I})$$



# Mapping from Latent Process to Observed

$$\begin{bmatrix} W & 0 & 0 \\ 0 & W & 0 \\ 0 & 0 & W \end{bmatrix} \times \begin{bmatrix} \mathbf{x}_{1,:} \\ \mathbf{x}_{2,:} \\ \mathbf{x}_{3,:} \end{bmatrix} = \begin{bmatrix} W\mathbf{x}_{1,:} \\ W\mathbf{x}_{2,:} \\ W\mathbf{x}_{3,:} \end{bmatrix}$$

# Output Covariance

This leads to a covariance of the form

$$(\mathbf{I} \otimes \mathbf{W})(\mathbf{K} \otimes \mathbf{I})(\mathbf{I} \otimes \mathbf{W}^\top) + \mathbf{I}\sigma^2$$

Using  $(\mathbf{A} \otimes \mathbf{B})(\mathbf{C} \otimes \mathbf{D}) = \mathbf{AC} \otimes \mathbf{BD}$  This leads to

$$\mathbf{K} \otimes \mathbf{WW}^\top + \mathbf{I}\sigma^2$$

or

$$\mathbf{y} \sim \mathcal{N}(0, \mathbf{WW}^\top \otimes \mathbf{K} + \mathbf{I}\sigma^2)$$

# Kernels for Vector Valued Outputs: A Review

Foundations and Trends<sup>®</sup> in  
Machine Learning  
Vol. 4, No. 3 (2011) 195–266  
© 2012 M. A. Álvarez, L. Rosasco and N. D. Lawrence  
DOI: 10.1561/22000000036



## **Kernels for Vector-Valued Functions: A Review**

By Mauricio A. Álvarez,  
Lorenzo Rosasco and Neil D. Lawrence

# Kronecker Structure GPs

- ▶ This Kronecker structure leads to several published models.

$$(\mathbf{K}(\mathbf{x}, \mathbf{x}'))_{d,d'} = k(\mathbf{x}, \mathbf{x}')k_T(d, d'),$$

where  $k$  has  $\mathbf{x}$  and  $k_T$  has  $n$  as inputs.

- ▶ Can think of multiple output covariance functions as covariances with augmented input.
- ▶ Alongside  $\mathbf{x}$  we also input the  $d$  associated with the *output* of interest.

# Separable Covariance Functions

- ▶ Taking  $\mathbf{B} = \mathbf{W}\mathbf{W}^\top$  we have a matrix expression across outputs.

$$\mathbf{K}(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}, \mathbf{x}')\mathbf{B},$$

where  $\mathbf{B}$  is a  $p \times p$  symmetric and positive semi-definite matrix.

- ▶  $\mathbf{B}$  is called the *coregionalization* matrix.
- ▶ We call this class of covariance functions *separable* due to their product structure.

# Sum of Separable Covariance Functions

- ▶ In the same spirit a more general class of kernels is given by

$$\mathbf{K}(\mathbf{x}, \mathbf{x}') = \sum_{j=1}^q k_j(\mathbf{x}, \mathbf{x}') \mathbf{B}_j.$$

- ▶ This can also be written as

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \sum_{j=1}^q \mathbf{B}_j \otimes k_j(\mathbf{X}, \mathbf{X}),$$

- ▶ This is like several Kalman filter-type models added together, but each one with a different set of latent functions.
- ▶ We call this class of kernels sum of separable kernels (SoS kernels).

- ▶ Use of GPs in Geostatistics is called kriging.
- ▶ These multi-output GPs pioneered in geostatistics: prediction over vector-valued output data is known as *cokriging*.
- ▶ The model in geostatistics is known as the *linear model of coregionalization* (LMC, Journel and Huijbregts (1978); Goovaerts (1997)).
- ▶ Most machine learning multitask models can be placed in the context of the LMC model.

# Weighted sum of Latent Functions

- ▶ In the linear model of coregionalization (LMC) outputs are expressed as linear combinations of independent random functions.
- ▶ In the LMC, each component  $f_d$  is expressed as a linear sum

$$f_d(\mathbf{x}) = \sum_{j=1}^q w_{d,j} u_j(\mathbf{x}).$$

where the latent functions are independent and have covariance functions  $k_j(\mathbf{x}, \mathbf{x}')$ .

- ▶ The processes  $\{f_j(\mathbf{x})\}_{j=1}^q$  are independent for  $q \neq j'$ .



# Kalman Filter Special Case

- ▶ The Kalman filter is an example of the LMC where  $u_i(\mathbf{x}) \rightarrow x_i(t)$ .
- ▶ I.e. we've moved from time input to a more general input space.
- ▶ In matrix notation:

1. Kalman filter

$$\mathbf{F} = \mathbf{W}\mathbf{X}$$

2. LMC

$$\mathbf{F} = \mathbf{W}\mathbf{U}$$

where the rows of these matrices  $\mathbf{F}$ ,  $\mathbf{X}$ ,  $\mathbf{U}$  each contain  $q$  samples from their corresponding functions at a different time (Kalman filter) or spatial location (LMC).

# Intrinsic Coregionalization Model

- ▶ If one covariance used for latent functions (like in Kalman filter).
- ▶ This is called the intrinsic coregionalization model (ICM, Goovaerts (1997)).
- ▶ The kernel matrix corresponding to a dataset  $\mathbf{X}$  takes the form

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{B} \otimes k(\mathbf{X}, \mathbf{X}).$$

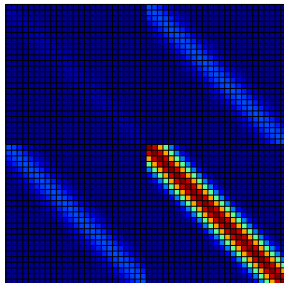
# Autokrigeability

- ▶ If outputs are noise-free, maximum likelihood is equivalent to independent fits of  $\mathbf{B}$  and  $k(\mathbf{x}, \mathbf{x}')$  (Helterbrand and Cressie, 1994).
- ▶ In geostatistics this is known as autokrigeability (Wackernagel, 2003).
- ▶ In multitask learning its the cancellation of intertask transfer (Bonilla et al., 2008).

# Intrinsic Coregionalization Model

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{w}\mathbf{w}^\top \otimes k(\mathbf{X}, \mathbf{X}).$$

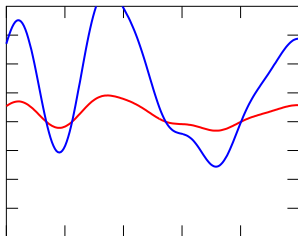
$$\mathbf{w} = \begin{bmatrix} 1 \\ 5 \end{bmatrix}$$
$$\mathbf{B} = \begin{bmatrix} 1 & 5 \\ 5 & 25 \end{bmatrix}$$



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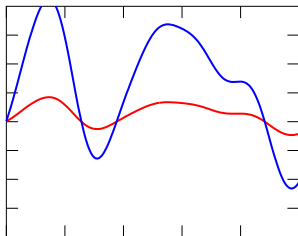
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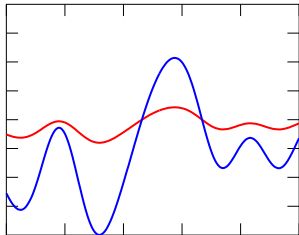
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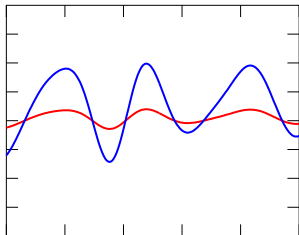
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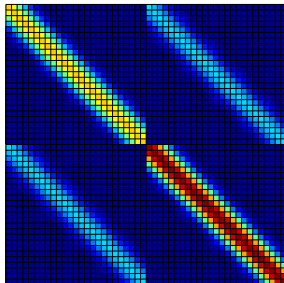




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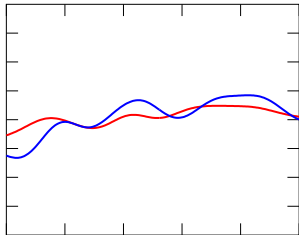
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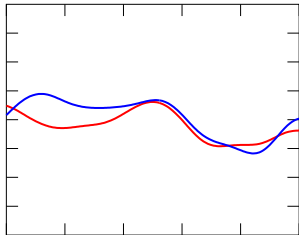
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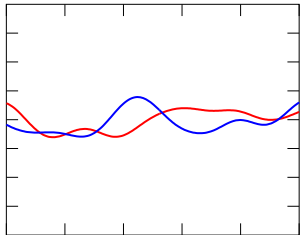
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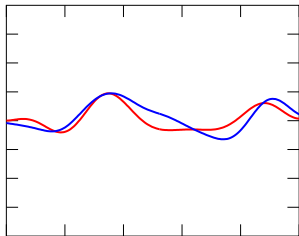
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# LMC Samples

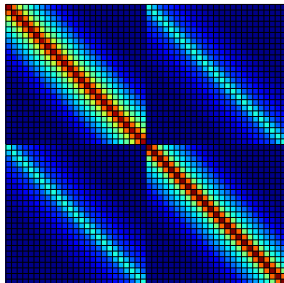
$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{B}_1 \otimes k_1(\mathbf{X}, \mathbf{X}) + \mathbf{B}_2 \otimes k_2(\mathbf{X}, \mathbf{X})$$

$$\mathbf{B}_1 = \begin{bmatrix} 1.4 & 0.5 \\ 0.5 & 1.2 \end{bmatrix}$$

$$\ell_1 = 1$$

$$\mathbf{B}_2 = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1.3 \end{bmatrix}$$

$$\ell_2 = 0.2$$



# LMC Samples

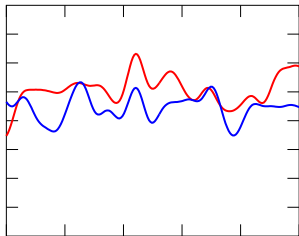
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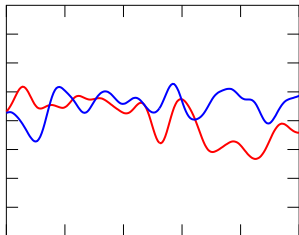
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# LMC Samples

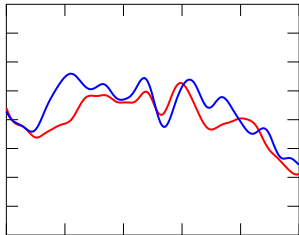
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$$\ell_2 = 0.2$$



# LMC Samples

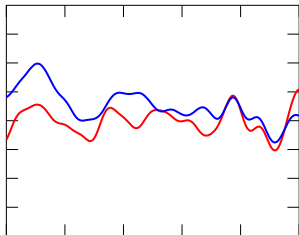
$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{B}_1 \otimes k_1(\mathbf{X}, \mathbf{X}) + \mathbf{B}_2 \otimes k_2(\mathbf{X}, \mathbf{X})$$

$$\mathbf{B}_1 = \begin{bmatrix} 1.4 & 0.5 \\ 0.5 & 1.2 \end{bmatrix}$$

$$\ell_1 = 1$$

$$\mathbf{B}_2 = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1.3 \end{bmatrix}$$

$$\ell_2 = 0.2$$



# LMC in Machine Learning and Statistics

- ▶ Used in machine learning for GPs for multivariate regression and in statistics for computer emulation of expensive multivariate computer codes.
- ▶ Imposes the correlation of the outputs explicitly through the set of coregionalization matrices.
- ▶ Setting  $\mathbf{B} = \mathbf{I}_p$  assumes outputs are conditionally independent given the parameters  $\boldsymbol{\theta}$ . (Minka and Picard, 1997; Lawrence and Platt, 2004; Yu et al., 2005).
- ▶ More recent approaches for multiple output modeling are different versions of the linear model of coregionalization.

# Semiparametric Latent Factor Model

- ▶ Coregionalization matrices are rank 1 Teh et al. (2005).  
rewrite equation (??) as

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \sum_{j=1}^q \mathbf{w}_{:,j} \mathbf{w}_{:,j}^{\top} \otimes k_j(\mathbf{X}, \mathbf{X}).$$

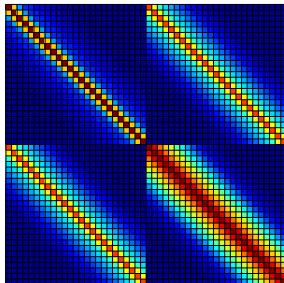
- ▶ Like the Kalman filter, but each latent function has a *different* covariance.
- ▶ Authors suggest using an exponentiated quadratic characteristic length-scale for each input dimension.

# Semiparametric Latent Factor Model Samples

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{w}_{:,1} \mathbf{w}_{:,1}^\top \otimes k_1(\mathbf{X}, \mathbf{X}) + \mathbf{w}_{:,2} \mathbf{w}_{:,2}^\top \otimes k_2(\mathbf{X}, \mathbf{X})$$

$$\mathbf{w}_1 = \begin{bmatrix} 0.5 \\ 1 \end{bmatrix}$$

$$\mathbf{w}_2 = \begin{bmatrix} 1 \\ 0.5 \end{bmatrix}$$

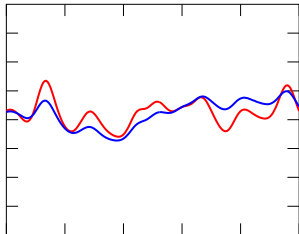


# Semiparametric Latent Factor Model Samples

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{w}_{:,1} \mathbf{w}_{:,1}^\top \otimes k_1(\mathbf{X}, \mathbf{X}) + \mathbf{w}_{:,2} \mathbf{w}_{:,2}^\top \otimes k_2(\mathbf{X}, \mathbf{X})$$

$$\mathbf{w}_1 = \begin{bmatrix} 0.5 \\ 1 \end{bmatrix}$$

$$\mathbf{w}_2 = \begin{bmatrix} 1 \\ 0.5 \end{bmatrix}$$

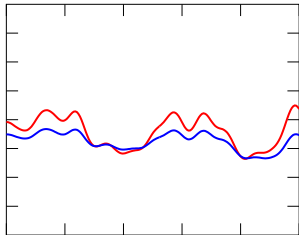


# Semiparametric Latent Factor Model Samples

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{w}_{:,1} \mathbf{w}_{:,1}^\top \otimes k_1(\mathbf{X}, \mathbf{X}) + \mathbf{w}_{:,2} \mathbf{w}_{:,2}^\top \otimes k_2(\mathbf{X}, \mathbf{X})$$

$$\mathbf{w}_1 = \begin{bmatrix} 0.5 \\ 1 \end{bmatrix}$$

$$\mathbf{w}_2 = \begin{bmatrix} 1 \\ 0.5 \end{bmatrix}$$

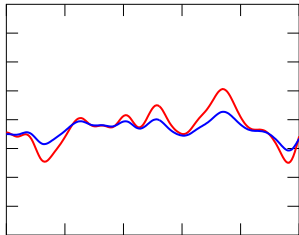


# Semiparametric Latent Factor Model Samples

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{w}_{:,1} \mathbf{w}_{:,1}^\top \otimes k_1(\mathbf{X}, \mathbf{X}) + \mathbf{w}_{:,2} \mathbf{w}_{:,2}^\top \otimes k_2(\mathbf{X}, \mathbf{X})$$

$$\mathbf{w}_1 = \begin{bmatrix} 0.5 \\ 1 \end{bmatrix}$$

$$\mathbf{w}_2 = \begin{bmatrix} 1 \\ 0.5 \end{bmatrix}$$



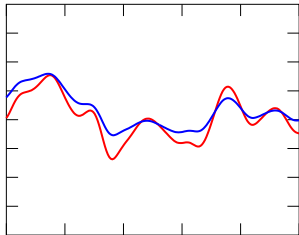


# Semiparametric Latent Factor Model Samples

$$\mathbf{K}(\mathbf{X}, \mathbf{X}) = \mathbf{w}_{:,1} \mathbf{w}_{:,1}^\top \otimes k_1(\mathbf{X}, \mathbf{X}) + \mathbf{w}_{:,2} \mathbf{w}_{:,2}^\top \otimes k_2(\mathbf{X}, \mathbf{X})$$

$$\mathbf{w}_1 = \begin{bmatrix} 0.5 \\ 1 \end{bmatrix}$$

$$\mathbf{w}_2 = \begin{bmatrix} 1 \\ 0.5 \end{bmatrix}$$



# Gaussian processes for Multi-task, Multi-output and Multi-class

- ▶ Bonilla et al. (2008) suggest ICM for multitask learning.
- ▶ Use a PPCA form for  $\mathbf{B}$ : similar to our Kalman filter example.
- ▶ Refer to the autokrigeability effect as the cancellation of inter-task transfer.
- ▶ Also discuss the similarities between the multi-task GP and the ICM, and its relationship to the SLFM and the LMC.

# Multitask Classification

- ▶ Mostly restricted to the case where the outputs are conditionally independent given the hyperparameters  $\phi$  (Minka and Picard, 1997; Williams and Barber, 1998; Lawrence and Platt, 2004; Seeger and Jordan, 2004; Yu et al., 2005; Rasmussen and Williams, 2006).
- ▶ Intrinsic coregionalization model has been used in the multiclass scenario. Skolidis and Sanguinetti (2011) use the intrinsic coregionalization model for classification, by introducing a probit noise model as the likelihood.
- ▶ Posterior distribution is no longer analytically tractable: approximate inference is required.

# Computer Emulation

- ▶ A statistical model used as a surrogate for a computationally expensive computer model.
- ▶ Higdon et al. (2008) use the linear model of coregionalization to model images representing the evolution of the implosion of steel cylinders.
- ▶ In Conti and O'Hagan (2009) use the ICM to model a vegetation model: called the Sheffield Dynamic Global Vegetation Model (Woodward et al., 1998).

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