An introduction to Gaussian Processes

Richard Wilkinson

School of Maths and Statistics University of Sheffield

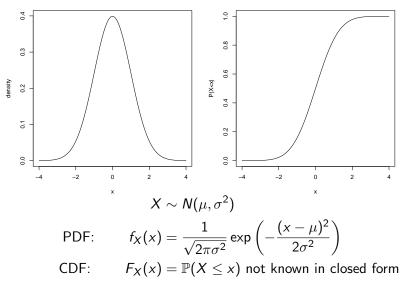
GP summer school September 2019

Introduction

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PDF of a N(0,1) random variable

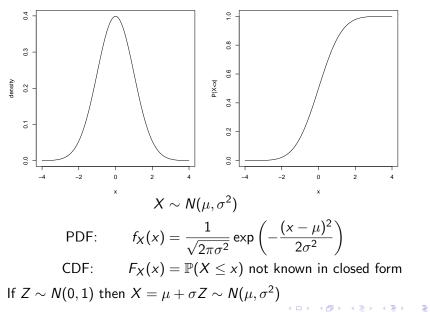
CDF of a N(0,1) random variable



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- Square-loss functions lead to procedures that have a Gaussian probabilistic interpretation eg Fit model $f_{\beta}(x)$ to data y by mimizing $\sum (y_i f_{\beta}(x_i))^2$ is equivalent to maximum likelihood estimation under the assumption that $y = f_{\beta}(x) + \epsilon$ where $\epsilon \sim N(0, \sigma^2)$.

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Suppose $X \in \mathbb{R}^d$ has a multivariate Gaussian distribution with

• mean vector $\mu \in \mathbb{R}^d$

• covariance matrix
$$\Sigma \in \mathbb{R}^{d \times d}$$

Write

$$X \sim N_d(\mu, \Sigma)$$

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Bivariate Gaussian: d=2

$$X = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \qquad \mu = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} \qquad \Sigma = \begin{pmatrix} \sigma_1^2 & \rho_{12}\sigma_1\sigma_2 \\ \rho_{21}\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$$

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ho_{ij}\sigma_i\sigma_j \quad \mathsf{C}or(X_i, X_j) =
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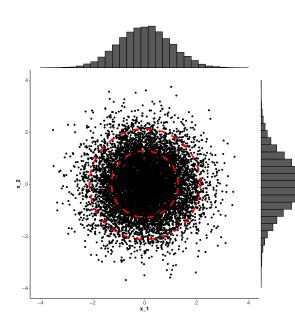
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pdf:
$$f(x \mid \mu, \Sigma) = |\Sigma|^{-\frac{1}{2}} (2\pi)^{-\frac{d}{2}} \exp\left(-\frac{1}{2}(x-\mu)^{\top} \Sigma^{-1}(x-\mu)\right)$$

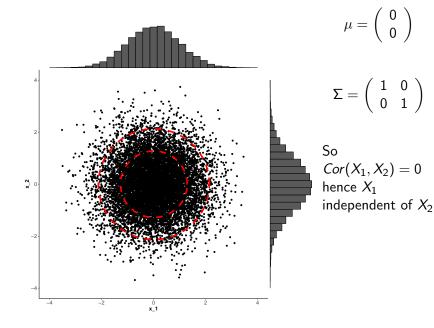
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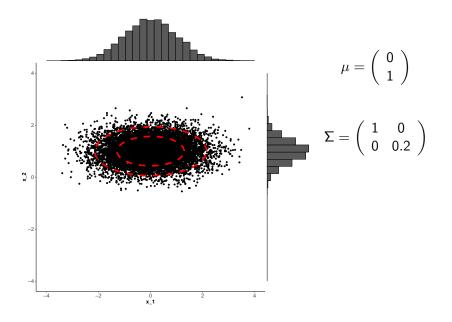
 $\mu = \left(\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right)$

 $\Sigma = \left(\begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array} \right)$

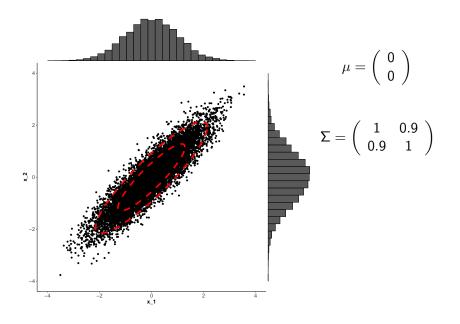
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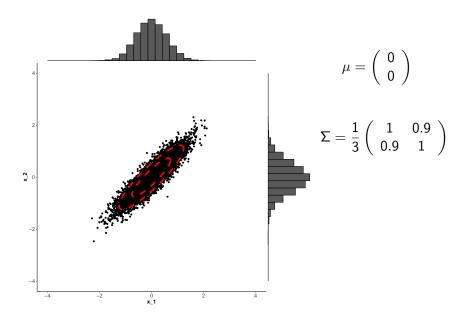
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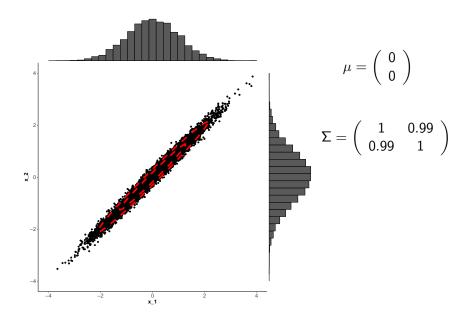
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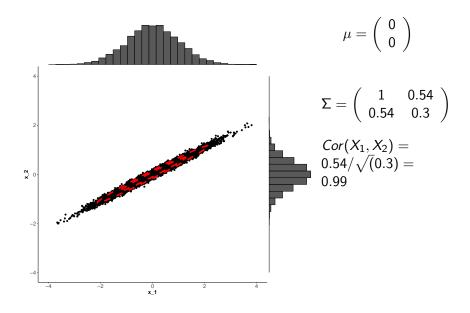
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More pictures

Consider d = 5 with

$$\Sigma = \left(egin{array}{cccccc} 1 & 0.9 & 0.8 & 0.7 & 0.6 \ 0.9 & 1 & 0.9 & 0.8 & 0.7 \ 0.8 & 0.9 & 1 & 0.9 & 0.8 \ 0.7 & 0.8 & 0.9 & 1 & 0.9 \ 0.6 & 0.7 & 0.8 & 0.9 & 1 \end{array}
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It's hard to visualise in dimensions > 3, so let's stack points next to each other.

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More pictures

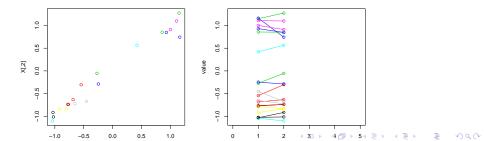
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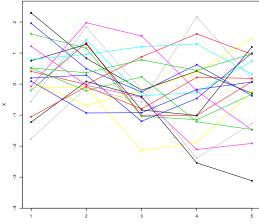
So for 2d instead of

we have



d=5

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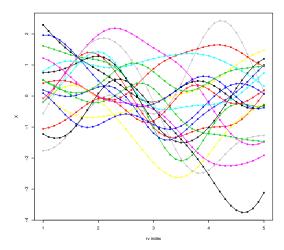


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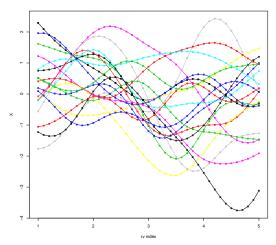
d=50

	/ 1	0.99	0.98	0.97	0.96	\
$\Sigma =$	0.99	1	0.99	0.98	0.97)
	0.98	0.99	1	0.99	0.98	
	0.97	0.98	0.99	1	0.99	
	0.96	0.97	0.98	0.99	1	
						.]
	\ ·					• /



d=50

Σ =	$\begin{pmatrix} 1 \\ 0.99 \end{pmatrix}$	0.99	0.98 0.99	0.97 0.98	0.96 0.97)
	0.98	0.99	1 0.99	0.99	0.98	
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	(:	:	:	:	:	:)



A stochastic process is a collection of random variables indexed by some variable $x \in \mathcal{X}$

 $f = \{f(x) : x \in \mathcal{X}\}$

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Thankfully we only need consider the finite dimensional distributions (FDDs), i.e., for all x_1, \ldots, x_n and for all $n \in \mathbb{N}$

$$\mathbb{P}(f(x_1) \leq y_1, \ldots, f(x_n) \leq y_n)$$

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Why would we want to use this very restricted class of model?

Gaussian **distributions** have several properties that make them easy to work with:

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Conversely, any matrix $\boldsymbol{\Sigma}$ which is positive semi-definite is a valid covariance matrix:

If
$$Z \sim N_d(0_d, I_d)$$
 then $X = \mu + \Sigma^{\frac{1}{2}} Z \sim N_d(\mu, \Sigma)$.
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• Gives one way of generating multivariate Gaussians.

Property 2: Conditional distributions are still Gaussian

Suppose

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where

$$\mu = \left(\begin{array}{c} \mu_1 \\ \mu_2 \end{array}\right) \qquad \Sigma = \left(\begin{array}{cc} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{array}\right)$$

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Then

$$X_2 \mid X_1 = x_1 \sim N\left(\mu_2 + \Sigma_{21}\Sigma_{11}^{-1}(x_1 - \mu_1), \Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12}
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$$\pi(x_2|x_1) = \frac{\pi(x_1, x_2)}{\pi(x_1)} \propto \pi(x_1, x_2)$$

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$$\pi(x_2|x_1) = \frac{\pi(x_1, x_2)}{\pi(x_1)} \propto \pi(x_1, x_2)$$
$$\propto \exp\left(-\frac{1}{2}(x-\mu)^{\top} \Sigma^{-1}(x-\mu)\right)$$
$$= \exp\left(-\frac{1}{2}\left[\left(\begin{pmatrix}x_1\\x_2\end{pmatrix} - \begin{pmatrix}\mu_1\\\mu_2\end{pmatrix}\right)^{\top}\begin{pmatrix}Q_{11} & Q_{12}\\Q_{21} & Q_{22}\end{pmatrix}\cdots\right]$$

where

$$\Sigma^{-1} := Q := \left(egin{array}{cc} Q_{11} & Q_{12} \ Q_{21} & Q_{22} \end{array}
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$$\propto \exp\left(-\frac{1}{2}\left[(x_{2}-\mu_{2})^{\top}Q_{22}(x_{2}-\mu_{2}) + 2(x_{2}-\mu_{2})^{\top}Q_{21}(x_{1}-\mu_{1})\right]\right)$$

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$$\propto \exp\left(-\frac{1}{2}\left[x_{2}^{\top}Q_{22}x_{2}-2x_{2}^{\top}(Q_{22}\mu_{2}+Q_{21}(x_{1}-\mu_{1}))\right]\right)$$

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$$\propto \exp\left(-\frac{1}{2}\left(x_{2}-Q_{22}^{-1}(Q_{22}\mu_{2}+Q_{21}(x_{1}-\mu_{1}))\right)^{\top}Q_{22}(x_{2}-\ldots)\right)$$

$$\pi(x_{2}|x_{1}) \propto \exp\left(-\frac{1}{2}\left[(x_{2}-\mu_{2})^{\top}Q_{22}(x_{2}-\mu_{2})+2(x_{2}-\mu_{2})^{\top}Q_{21}(x_{1}-\mu_{1})\right]\right)$$
$$\propto \exp\left(-\frac{1}{2}\left[x_{2}^{\top}Q_{22}x_{2}-2x_{2}^{\top}(Q_{22}\mu_{2}+Q_{21}(x_{1}-\mu_{1}))\right]\right)$$
$$\propto \exp\left(-\frac{1}{2}\left(x_{2}-Q_{22}^{-1}(Q_{22}\mu_{2}+Q_{21}(x_{1}-\mu_{1}))\right)^{\top}Q_{22}(x_{2}-\ldots)\right)$$
So

$$X_2|X_1 = x_1 \sim N(\mu_2 + Q_{22}^{-1}Q_{21}(x_1 - \mu_1), Q_{22})$$

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$$\pi(x_{2}|x_{1}) \propto \exp\left(-\frac{1}{2}\left[(x_{2}-\mu_{2})^{\top}Q_{22}(x_{2}-\mu_{2})+2(x_{2}-\mu_{2})^{\top}Q_{21}(x_{1}-\mu_{1})\right]\right)$$
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So
$$X_{2}|X_{1}=x_{1} \sim N(\mu_{2}+Q_{21}^{-1}Q_{22}(x_{1}-\mu_{1}),Q_{22})$$

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 $\lambda_2 | \lambda_1 = x_1 \sim N(\mu_2 + Q_{22} Q_{21}(x_1 - \mu_1), Q_{22})$

A simple matrix inversion lemma gives

$$\begin{split} Q_{22}^{-1} &= \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12} \\ \text{and} \, Q_{22}^{-1} \, Q_{21} &= \Sigma_{21} \Sigma_{11}^{-1} \end{split}$$

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A simple matrix inversion lemma gives

$$Q_{22}^{-1} = \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12}$$

and $Q_{22}^{-1} Q_{21} = \Sigma_{21} \Sigma_{11}^{-1}$

giving

$$X_2|X_1 = x_1 \sim N\left(\mu_2 + \Sigma_{21}\Sigma_{11}^{-1}(x_1 - \mu_1), \Sigma_{22} - \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12}\right)$$

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Conditional updates of Gaussian processes

So suppose f is a Gaussian process, then

 $f(x_1),\ldots,f(x_n),f(x) \sim N(\mu,\Sigma)$



Conditional updates of Gaussian processes So suppose f is a Gaussian process, then

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If we observe its value at x_1, \ldots, x_n then

$$f(x)|f(x_1),\ldots,f(x_n)\sim N(\mu^*,\sigma^*)$$

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where μ^* and σ^* are as on the previous slide.

Conditional updates of Gaussian processes

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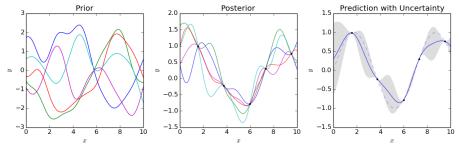
$$f(x_1),\ldots,f(x_n),f(x)\sim N(\mu,\Sigma)$$

If we observe its value at x_1, \ldots, x_n then

$$f(x)|f(x_1),\ldots,f(x_n)\sim N(\mu^*,\sigma^*)$$

where μ^* and σ^* are as on the previous slide.

Note that we still believe f is a GP even though we've observed its value at a number of locations.



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The GP class of models is closed under various operations.

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• Closed under addition

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$$\mathbf{D}=(f(x_1),\ldots,f(x_n))$$

then

$$f|D \sim GP$$

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but with updated mean and covariance functions.

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but with updated mean and covariance functions.

Closed under any linear operator. If f ~ GP(m(·), k(·, ·)), then if L is a linear operator

$$\mathcal{L} \circ f \sim GP(\mathcal{L} \circ m, \mathcal{L}^2 \circ k)$$

e.g.
$$\frac{df}{dx}$$
, $\int f(x)dx$, Af are all GPs

How do we determine the mean $\mathbb{E}(f(x))$ and covariance $\mathbb{C}ov(f(x), f(x'))$?

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• We can use any mean function we want:

$$m(x) = \mathbb{E}(f(x))$$

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Most popular choices are m(x) = 0 or m(x) = a for all x, or m(x) = a + bx

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$$m(x) = \mathbb{E}(f(x))$$

Most popular choices are m(x) = 0 or m(x) = a for all x, or m(x) = a + bx

• If mean is a linear combination of known regressor functions, e.g.,

$$m(x) = \beta h(x)$$
 for known $h(x)$

and $\beta \sim N(\cdot, \cdot)$ is given a normal prior (including $\pi(\beta) \propto 1$), then $f|D, \beta \sim GP$ and

$$f|D \sim GP$$

with slightly modified mean and variance formulas.

Covariance functions

• We usually use a covariance function that is a function of distance between the locations

$$k(x,x') = \mathbb{C}\mathrm{ov}(f(x),f(x')),$$

which has to be positive semi-definite, i.e., lead to valid covariance matrices.

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Covariance functions

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This can be problematic (see Nicolas' talk)

• If

$$k(x,x') = \sigma^2 c(x,x')$$

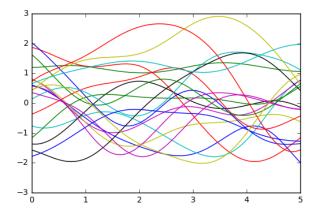
and we give σ^2 an inverse gamma prior (including $\pi(\sigma^2) \propto 1/\sigma^2$) then $f|D, \sigma^2 \sim GP$ and

$$f|D \sim t$$
-process

with n - p degrees of freedom. In practice, for reasonable *n*, this is indistinguishable from a GP.

 $\mathsf{RBF}/\mathsf{Squared}\text{-exponential}/\mathsf{exponentiated}\ \mathsf{quadratic}$

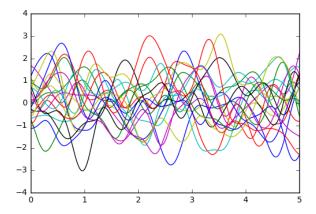
$$k(x,x') = \exp\left(-\frac{1}{2}(x-x')^2\right)$$



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RBF/Squared-exponential/exponentiated quadratic

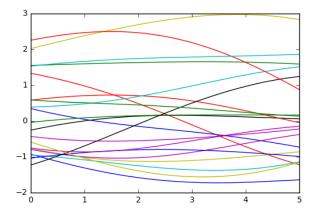
$$k(x, x') = \exp\left(-\frac{1}{2}\frac{(x-x')^2}{0.25^2}\right)$$



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RBF/Squared-exponential/exponentiated quadratic

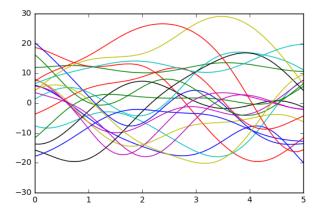
$$k(x, x') = \exp\left(-\frac{1}{2}\frac{(x - x')^2}{4^2}\right)$$



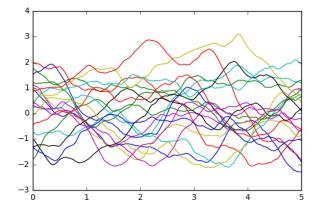
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RBF/Squared-exponential/exponentiated quadratic

$$k(x, x') = 100 \exp\left(-\frac{1}{2}(x - x')^2\right)$$



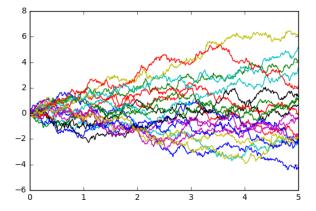
Matern 3/2 $k(x,x') \sim (1+|x-x'|) \exp\left(-|x-x'|
ight)$



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Brownian motion

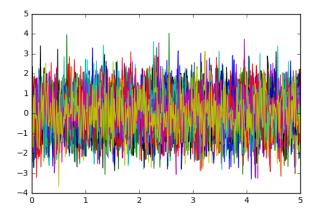
$$k(x,x') = \min(x,x')$$



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White noise

$$k(x,x') = egin{cases} 1 & ext{if } x = x' \ 0 & ext{otherwise} \end{cases}$$



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The GP inherits its properties primarily from the covariance function k.

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- Smoothness
- Differentiability
- Variance

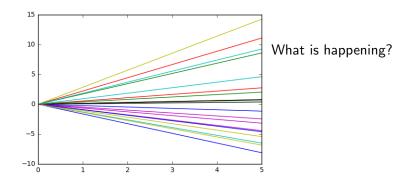
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- A final example

$$k(x,x') = x^{\top}x'$$

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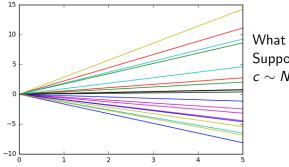
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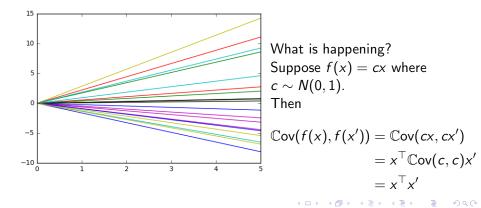
What is happening? Suppose f(x) = cx where $c \sim N(0, 1)$.

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The GP inherits its properties primarily from the covariance function k.

- Smoothness
- Differentiability
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$$k(x,x') = x^{\top}x'$$



So suppose f is a Gaussian process, then

$$f(x_1),\ldots,f(x_n),f(x)\sim N(0,\Sigma)$$

where

$$\Sigma = \begin{pmatrix} & & & & k(x_1, x) \\ & & & & k(x_2, x) \\ & & & & \vdots \\ & & & & k(x_n, x) \\ & & & & k(x, x_1) & k(x, x_2) & \dots & k(x, x_n) & k(x, x) \end{pmatrix}$$

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where $K_{ij} = k(x_i, x_j)$ is the Gram/kernel matrix.

Then

$$f(x)|f(x_1),\ldots,f(x_n) \sim N(m(x),c(x))$$

where

$$m(x) = k(x)(K + \sigma^2 I)^{-1}y$$

with

$$k(x) = (k(x, x_1) \ k(x, x_2) \ \dots \ k(x, x_n)) \in \mathbb{R}^{1 \times n}$$

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Cf

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We can also view GPs as a non-parametric extension to linear regression.

k determines the space of functions that sample paths live in.

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Suppose we're given data $\{(x_i, y_i)_{i=1}^n\}$.

Linear regression $y = x^{\top}\beta + \epsilon$ can be written solely in terms of inner products $x^{\top}x$.

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$$\hat{\beta} = \arg\min ||y - X\beta||_2^2 + \sigma^2 ||\beta||_2^2$$
$$= (X^\top X + \sigma^2 I)^{-1} X^\top y$$

where
$$X = \begin{pmatrix} x_{1}^{\top} \\ x_{2}^{\top} \\ \vdots \\ x_{n}^{\top} \end{pmatrix}$$

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where
$$X = \begin{pmatrix} x_1^T \\ x_2^T \\ \vdots \\ x_n^T \end{pmatrix}$$

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where
$$X = \begin{pmatrix} x_1^\top \\ x_2^\top \\ \vdots \\ x_n^\top \end{pmatrix}$$

At first the dual form

$$\hat{\beta} = X^{\top} (XX^{\top} + \sigma^2 I)^{-1} y$$

looks harder to compute than the usual

$$\hat{\beta} = (X^{\top}X + \sigma^2 I)^{-1}X^{\top}y$$

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• $X^{\top}X$ is $p \times p$ p = number of features/parameters • XX^{\top} is $n \times n$ n is the number of data points At first the dual form

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But the dual form only uses inner products.

$$XX^{\top} = \begin{pmatrix} x_1^{\top} \\ \vdots \\ x_n^{\top} \end{pmatrix} (x_1 \dots x_n) = \begin{pmatrix} x_1^{\top} x_1 & \dots & x_1^{\top} x_n \\ \vdots & & \\ x_n^{\top} x_1 & \dots & x_n^{\top} x_n \end{pmatrix} = K$$

— This is useful!

Prediction

The best prediction of y at a new location x' is

$$\begin{split} \hat{y}' &= x'^{\top} \hat{\beta} \\ &= x'^{\top} X^{\top} (XX^{\top} + \sigma^2 I)^{-1} y \\ &= k(x') (K + \sigma^2 I)^{-1} y \end{split}$$

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where $k(x') := (x'^{\top}x_1, \dots, x'^{\top}x_n)$ and $K_{ij} := x_i^{\top}x_j$

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where $k(x') := (x'^{\top}x_1, \dots, x'^{\top}x_n)$ and $K_{ij} := x_i^{\top}x_j$ K and k(x) are kernel matrices

• every element is an inner product btwn 2 points.

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where $k(x') := (x'^{\top}x_1, \dots, x'^{\top}x_n)$ and $K_{ij} := x_i^{\top}x_j$ K and k(x) are kernel matrices

every element is an inner product btwn 2 points.
 Note this is exactly the GP conditional mean we derived before.

$$m(x) = t(x)(K + \sigma^2 I)^{-1}y$$

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• linear regression and GP regression are equivalent when $k(x, x') = x^{\top}x'$.

• We know that we can replace x by a feature vector in linear regression, e.g., $\phi(x)=(1 \; x \; x^2)$ etc. Then

$$K_{ij} = \phi(x_i)^\top \phi(x_j)$$
 etc

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• For some sets of features, the inner product is equivalent to evaluating a kernel function

$$\phi(x)^{\top}\phi(x') \equiv k(x,x')$$

where

$$k: \mathcal{X} \times \mathcal{X} \to \mathbb{R}$$

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is a semi-positive definite function.

• For some sets of features, the inner product is equivalent to evaluating a kernel function

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$$k: \mathcal{X} \times \mathcal{X} \to \mathbb{R}$$

is a semi-positive definite function. **Example:** If (modulo some detail)

$$\phi(x) = \left(e^{-\frac{(x-c_1)^2}{2\lambda^2}}, \dots, e^{-\frac{(x-c_N)^2}{2\lambda^2}}\right)$$

then as $N
ightarrow \infty$ then

$$\phi(x)^{\top}\phi(x) = \exp\left(-\frac{(x-x')^2}{2\lambda^2}\right)$$

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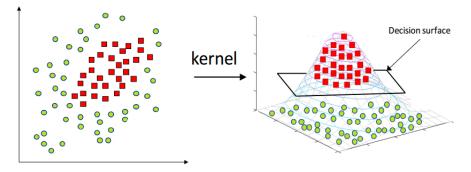
then as $N o \infty$ then

$$\phi(x)^{\top}\phi(x) = \exp\left(-\frac{(x-x')^2}{2\lambda^2}\right)$$

 We can use an infinite dimensional feature vector φ(x), and because linear regression can be done solely in terms of inner-products (inverting a n × n matrix in the dual form) we never need evaluate the feature vector, only the kernel.

Kernel trick:

lift x into feature space by replacing inner products $x^{\top}x'$ by k(x, x')



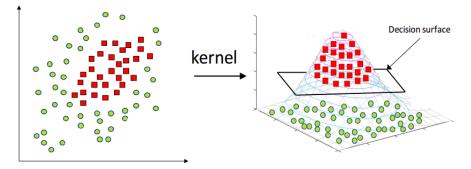
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Kernel trick:

lift x into feature space by replacing inner products $x^{\top}x'$ by k(x, x')



Kernel regression/non-parametric regression/GP regression all closely related:

$$\hat{y}' = m(x') = \sum_{i=1}^{n} \alpha_i k(x, x_i)$$

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Although reality may not lie in the RKHS defined by k, this space is much richer than any parametric regression model (and can be dense in some sets of continuous bounded functions), and is thus more likely to contain an element close to the true functional form than any class of models that contains only a finite number of features.

This is the motivation for non-parametric methods.

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Why use Gaussian processes as non-parametric models?

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One answer might come from Bayes linear methods². If we only knew the expectation and variance of some random variables, X and Y, then how should we best do statistics?

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It has been shown, using coherency arguments, or geometric arguments, or..., that the best second-order inference we can do to update our beliefs about X given Y is

$$\mathbb{E}(X|Y) = \mathbb{E}(X) + \mathbb{C}ov(X,Y)\mathbb{V}ar(Y)^{-1}(Y - \mathbb{E}(Y))$$

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See also kernel Bayes and kriging/BLUP.

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Why use GPs? Answer 4: Uncertainty estimates from emulators

We often think of our prediction as consisting of two parts

- point estimate
- uncertainty in that estimate

That GPs come equipped with the uncertainty in their prediction is seen as one of their main advantages.

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Warning: the uncertainty estimates from a GP can be flawed. Note that given data $D = \{X, y\}$

$$\mathbb{V}\operatorname{ar}(f(x)|X,y) = k(x,x) - k(x,X)k(X,X)^{-1}k(X,x)$$

so that the posterior variance of f(x) does not depend upon y!

The variance estimates are particularly sensitive to the hyper-parameter estimates.

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- We pick a covariance function from a small set, based usually on differentiability considerations.
- Possibly try a few (plus combinations of a few) covariance functions, and attempt to make a good choice using some sort of empirical evaluation.
- Covariance functions often contain hyper-parameters. E.g
 - RBF kernel

$$k(x, x') = \sigma^2 \exp\left(-\frac{1}{2}\frac{(x-x')^2}{\lambda^2}\right)$$

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Estimate these using some standard procedure (maximum likelihood, cross-validation, Bayes etc)

Gelman et al. 2017

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E.g. consider a zero mean GP on [0,1] with covariance function

$$k(x, x') = \sigma^2 \exp(-\kappa^2 |x - x|)$$

We can consistently estimate $\sigma^2 \kappa$, but not σ^2 or κ , even as $n \to \infty$.

Problems with hyper-parameter optimization

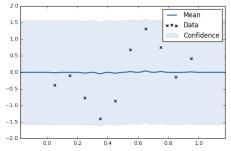
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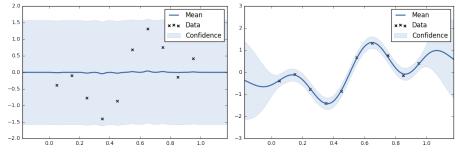
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We often work around these problems by running the optimizer multiple times from random start points, using prior distributions, constraining or fixing hyper-parameters, or adding white noise.