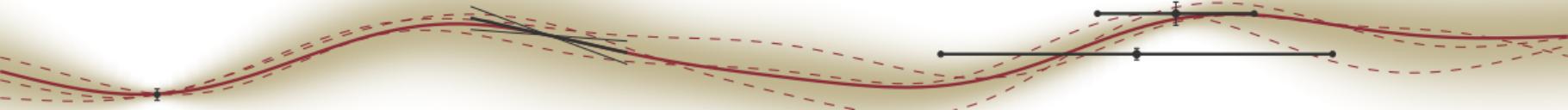


# ODE Solvers as Gauss-Markov Regression: An Overview

Filip Tronarp





Initial value problem:

$$\dot{y}^*(t) = f(y^*(t)), \quad y^*(0) = y_0, \quad t \in [0, T] \quad (1)$$

## Problem

- Grid:  $0 = t_0 < t_1 < \dots < t_N = T$
- Evaluations:  $f(\cdot)$

Find approximation:  $\hat{y} \approx y^*$



# A probabilistic formulation

At a glance

Probabilistic formulation:

- ♦ Prior:  $y \sim \mathcal{GP}$
- ♦ Initial data:  $y(0) = y^*(0)$
- ♦ Data:  $\dot{y}(t) = f(y(t))$  for  $t = t_0, t_1, \dots, t_N$
- ♦ Bayes' rule

**Voilá!**



# State-space realisable priors

Convenient and canonical priors

Prior:

$$dy^{(\nu)}(t) = \sum_{m=0}^{\nu} A_m y^{(m)}(t) dt + \sqrt{\kappa} \sigma(t) dw(t) \quad (2)$$

Usually  $\nu$ -times integrated Wiener process:<sup>1</sup>

$$dy^{(\nu)}(t) = \sqrt{\kappa} dw(t) \quad (3)$$

Corresponds to Taylor polynomial + perturbation:

$$y(t) = \sum_{m=0}^{\nu} y^{(m)}(0) \frac{t^m}{m!} + \sqrt{\kappa} \int_0^t \frac{(t-\tau)^{\nu}}{\nu!} dw(\tau)$$

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<sup>1</sup>A probabilistic model for the numerical solution of initial value problems. M. Schober, S. Särkkä, P. Hennig. Statistics and Computing, 2019.



# State-space realisable priors

## State-space realisations

For instance:

$$x^*(t) = \begin{pmatrix} y^{(\nu)*} & y^{(\nu-1)*} & \dots & y^{(0)*} \end{pmatrix}$$

State-space realisation:

$$dx(t) = Ax(t) dt + B\sqrt{\kappa}\sigma(t) dw(t), \quad (4a)$$

$$y^{(m)}(t) = E_m x(t). \quad (4b)$$

- $x(t)$  is a Gauss–Markov process
- $y$  and its derivatives are linear transforms of  $x$ .



# State-space realisable priors

The Gauss–Markov property

$x$  is Markov:

$$p(x(t_{0:N})) = p(x(t_0)) \prod_{n=1}^N p(x(t_n) | x(t_{n-1})) \quad \text{for } t_0 < t_1 < \dots < t_N. \quad (5)$$

In our case:

$$p(x(t) | x(u)) = \mathcal{N}\left(x(t); \Phi(t, u)x(u), \kappa Q(t, u)\right) \quad (6)$$

Parameters:

$$\Phi(t, u) = e^{A(t-u)}, \quad (7a)$$

$$Q(t, u) = \int_u^t \Phi(t, \tau) B \sigma(\tau) \sigma^*(\tau) B^* \Phi^*(t, \tau) d\tau. \quad (7b)$$



# ODE solvers as Non-linear Gauss–Markov regression

The inference problem

Non-linear Gauss–Markov regression problem:<sup>2</sup>

$$x(t_n) | x(t_{n-1}) \sim \mathcal{N}\left(\Phi(t_n, t_{n-1})x(t_{n-1}), \kappa Q(t_n, t_{n-1})\right), \quad (8a)$$

$$0 = z(x(t_n)) = E_1 x(t_n) - f(E_0 x(t_n)) = y^{(1)}(t_n) - f(y(t_n)), \quad n = 1, \dots, N. \quad (8b)$$

- Initial value  $x_0$  set to exact value via auto-diff.<sup>3</sup>
- $\kappa$  can be used to calibrate the numerical uncertainty.<sup>4 5 6</sup>

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<sup>2</sup>Probabilistic solutions to ordinary differential equations as nonlinear Bayesian filtering: a new perspective. F. Tronarp, H. Kersting, S Särkkä, P Hennig. Statistics and Computing, 2019.

<sup>3</sup>Stable implementation of probabilistic ODE solvers. N Krämer, P. Hennig. arXiv:2012.10106, 2020.

<sup>4</sup>Probabilistic solutions to ordinary differential equations as nonlinear Bayesian filtering: a new perspective. F. Tronarp, H. Kersting, S Särkkä, P Hennig. Statistics and Computing, 2019.

<sup>5</sup>A probabilistic model for the numerical solution of initial value problems. M. Schober, S. Särkkä, P. Hennig. Statistics and Computing, 2019.

<sup>6</sup>Calibrated adaptive probabilistic ODE solvers. N. Bosch, P. Hennig, F. Tronarp. AISTATS, 2021.



# Practical inference strategies

Exact inference: Linear problems

Affine vector field:

$$f(y) = L(t)y + b(t). \quad (9)$$

Affine measurements:

$$C(t) = E_1 - L(t)E_0, \quad (10a)$$

$$z(x(t_n)) = E_1x(t_n) - f(E_0x(t_n)) = C(t_n)x(t_n) - b(t_n). \quad (10b)$$

Solution: Kalman filtering and Rauch–Tung–Striebel smoothing.<sup>7</sup>

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<sup>7</sup>Bayesian filtering and smoothing. S. Särkkä. Cambridge University Press, 2013.



# Practical inference strategies

Exact inference: the Kalman filter

Posterior marginal for data up to time  $t_n$ :  $p(x(t_n) \mid z(x(t_{0:n})) = 0) = \mathcal{N}\left(x(t_n); \mu(t_n), \kappa\Sigma(t_n)\right)$

## The Kalman filter

Predict:

$$\mu(t_n^-) = \Phi(t_n, t_{n-1})\mu(t_{n-1}), \quad (11a)$$

$$\Sigma(t_n^-) = \Phi(t_n, t_{n-1})\Sigma(t_{n-1})\Phi^*(t_n, t_{n-1}) + Q(t_n, t_{n-1}). \quad (11b)$$

Update:

$$S(t_n) = C(t_n)\Sigma(t_n^-)C^*(t_n), \quad (12a)$$

$$K(t_n) = \Sigma(t_n^-)C^*(t_n)S^{-1}(t_n), \quad (12b)$$

$$\mu(t_n) = \mu(t_n^-) + K(t_n)\left(b(t_n) - C(t_n)\mu(t_n^-)\right), \quad (12c)$$

$$\Sigma(t_n) = \Sigma(t_n^-) - K(t_n)S(t_n)K^*(t_n). \quad (12d)$$



# Practical inference strategies

Exact inference: the Rauch–Tung–Striebel smoother

Posterior marginal for all data:

$$p(x(t_n) \mid z(x(t_{0:N})) = 0) = \mathcal{N}\left(x(t_n); \xi(t_n), \kappa \Lambda(t_n)\right)$$

## Rauch–Tung–Striebel smoother

Backwards prediction:

$$\xi(t_{n-1}) = G(t_{n-1}, t_n) \left( \xi(t_n) - \mu(t_n^-) \right), \quad (13a)$$

$$\Lambda(t_{n-1}) = G(t_{n-1}, t_n) \Lambda(t_n) G^*(t_{n-1}, t_n) + V(t_{n-1}, t_n), \quad (13b)$$

where

$$G(t_{n-1}, t_n) = \Sigma(t_{n-1}) \Phi^*(t_n, t_{n-1}) \Sigma^{-1}(t_n^-), \quad (14a)$$

$$V(t_{n-1}, t_n) = \Sigma(t_{n-1}) - G(t_{n-1}, t_n) \Sigma(t_n^-) G^*(t_{n-1}, t_n). \quad (14b)$$



# Practical inference strategies

Approximate inference: sequential methods

Successive linearisation:

- Zeroth order method (explicit):<sup>8</sup>

$$f(E_0 x(t)) \approx f(E_0 \mu(t_n^-)).$$

- First order method (semi-implicit):<sup>9</sup>

$$f(E_0 x(t_n)) \approx f(E_0 \mu(t_n^-)) + J_f(E_0 \mu(t_n^-)) E_0 (x(t_n) - \mu(t_n^-))$$

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<sup>8</sup>A probabilistic model for the numerical solution of initial value problems. M. Schober, S. Särkkä, P. Hennig. Statistics and Computing, 2019.

<sup>9</sup>Probabilistic solutions to ordinary differential equations as nonlinear Bayesian filtering: a new perspective. F. Tronarp, H. Kersting, S Särkkä, P Hennig. Statistics and Computing, 2019.



# Practical inference strategies

Approximate inference: maximum a posteriori estimation

$$\hat{x}(t_{1:N}) = \arg \min_{x(t_{1:N})} \frac{1}{2} \sum_{n=1}^N \|x(t_n) - \Phi(t_n, t_{n-1})x(t_{n-1})\|_{Q^{-1}(t_n, t_{n-1})}^2, \quad (15)$$

subject to  $z(x(t_n)) = 0, \quad n = 1, \dots, N.$

Equivalent to minimum norm interpolation in RKHS:<sup>10</sup>

$$\hat{y} = \arg \min_y \int_0^{t_N} \left| \left( y^{(\nu+1)}(t) - \sum_{m=0}^{\nu} A_m y^{(m)}(t) \right) \right|^2 \sigma^{-2}(t) dt,$$

subject to  $z(x(t_n)) = 0, \quad n = 1, \dots, N.$

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<sup>10</sup>Bayesian ODE solvers: the maximum a posteriori estimate. F. Tronarp, S. Särkkä, P. Hennig. Statistics and Computing, 2021.



Linear test equation:

$$\dot{y}(t) = \Lambda y(t).$$

### Definition: A-stability

A method  $\hat{y}$  using a constant step-size is A-stable if  $\hat{y}(t)$  is asymptotically whenever  $\Lambda$  has eigenvalues strictly in the left-half plane.

- Classical approach: analyse roots of discrete time process.
- Probabilistic approach: exploit systems theory results relating to stabilising control.



# Stability

## A-Stability

- Constant measurement matrix (semi-implicit):

$$C = E_1 - \Lambda E_0.$$

- Let  $\sigma(t) = \text{const}$ , implies model matrices  $\Phi$ ,  $Q$ , and  $C$  are all constant for constant step-size.

### Generative form

$$x(t_n) = \Phi x(t_{n-1}) + Q^{1/2} w(t_n), \quad (16a)$$

$$0 = C x(t_n). \quad (16b)$$



# Stability

## A-Stability

### Definition (Absolute stabilisability).

The pair  $[\Phi, Q^{1/2}]$  is completely stabilisable if  $w^* Q^{1/2} = 0$  and  $w^* \Phi = \eta w^*$  for some constant  $\eta$  implies either  $|\eta| < 1$  or  $w = 0$ .

### Definition (Absolute detectability).

The pair  $[\Phi, C]$  is completely detectable if  $[\Phi^*, C^*]$  is completely stabilisable.

### Theorem

The semi-implicit solver is exponentially (and therefore A-stable) if and only if the pair  $[\Phi, Q^{1/2}]$  and  $[\Phi, C]$  are complete stabilisable and detectable, respectively.

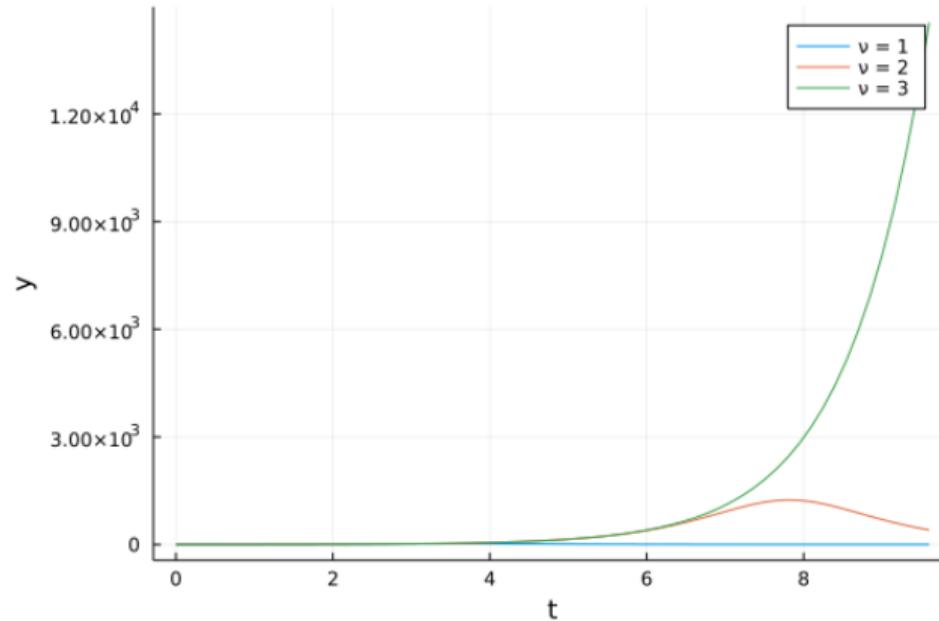


# Stability

Stable, all too stable?

- Complete detectability of  $[\Phi, C]$  is not a function of the real part of the eigenvalues of  $\Lambda$ !
- Let us use a  $\nu$ -times integrated Wiener process to solve:

$$\dot{y}^*(t) = y^*(t), \quad y^*(0) = 1.$$





# Convergence

A story with some gaps

Results for explicit methods:

- Matching methods associated with some priors to classical methods.<sup>11 12</sup>
- Local and global rates for a limited set of priors using “classical” convergence analysis.<sup>13</sup>

Results for semi-implicit methods:

- Only empirical so far.<sup>14 15</sup>

Results for MAP estimate:

- Quite nice result under mild assumptions using methods from scattered data approximation.<sup>16</sup>

**Contraction rates of the actual posterior has not been investigated at all?**

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<sup>11</sup>A probabilistic model for the numerical solution of initial value problems. M. Schober, S. Särkkä, P .Hennig. Statistics and Computing, 2019.

<sup>12</sup>Probabilistic ODE solvers with Runge-Kutta means. M. Schober, D. K. Duvenaud, P. Hennig. Neurips, 2014.

<sup>13</sup>Convergence rates of Gaussian ODE filters. H. Kersting, T. J. Sullivan, P. Hennig. Statistics and computing, 2020.

<sup>14</sup>Calibrated adaptive probabilistic ODE solvers. N. Bosch, P. Hennig, F. Tronarp. AISTATS, 2021.

<sup>15</sup>Stable implementation of probabilistic ODE solvers. N. Krämer, P. Hennig. arXiv:2012.10106, 2020.

<sup>16</sup>Bayesian ODE solvers: the maximum a posteriori estimate. F. Tronarp, S. Särkkä, P. Hennig. Statistics and Computing, 2021.



# Convergence

The MAP estimate: assumptions and consequences

Suppose:

- The prior is of the form:

$$dy^{(\nu)}(t) = \sum_{m=0}^{\nu} A_m y^{(m)}(t) dt + \sqrt{\kappa} dw(t). \quad (17)$$

- The vector field is smooth:  $f \in C^{\nu+1}$ .
- A unique solution  $y^*(t)$  exists up until  $T^* > t_N$ .

Then:

- RKHS is equivalent to  $H_2^{\nu+1}$ .
- The solution  $y^*(t)$  is in RKHS.
- The operator  $S_f[\varphi](t) = f(\varphi(t))$  is locally Lipschitz from  $B(0, \|y^*\|_{H_2^{\nu+1}}^2 + \varepsilon) \subset H_2^{\nu+1}$  onto  $H_2^\nu$ .<sup>17</sup>

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<sup>17</sup>Boundary Value Problems of Finite Elasticity: Local Theorems on Existence, Uniqueness, and Analytic Dependence on Data. T. Valent. Springer, 2013.



# Convergence

The MAP estimate: scattered data and nonlinear analysis

Integral form of estimate:

$$\hat{y}(t) = y(0) + \int_0^t \dot{\hat{y}}(\tau) d\tau = y(0) + \int_0^t f(\hat{y}(\tau)) d\tau + \int_0^t \dot{R}[\hat{y}; f](\tau) d\tau$$

Derivative of residual:

$$\dot{R}[\hat{y}; f](\tau) = \dot{\hat{y}}(\tau) - f(\hat{y}(\tau)) \quad (18)$$

Sobolev functions with many zeros are small:<sup>18</sup>

$$\left| \dot{R}_i[\hat{y}; f] \right|_{H_q^m} \leq c_2 h^{\nu-m-(1/2-1/q)_+} \left| \dot{R}_i[\hat{y}; f] \right|_{H_2^\nu}, \quad m \leq \nu - 1 \quad (19)$$

Lipschitz property and  $\hat{y}$  is smaller than  $y^*$ :

$$\left| \dot{R}_i[\hat{y}; f] \right|_{H_2^\nu} \leq \left\| \dot{R}_i[\hat{y}; f] - \dot{R}_i[y^*; f] \right\|_{H_2^\nu} \leq c_3^*(f) \left\| \hat{y} - y^* \right\|_{H_2^\nu} \leq 2c_3^*(f) \|y^*\|_{H_2^\nu} \quad (20)$$

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<sup>18</sup>An extension of a bound for functions in Sobolev spaces, with applications to (m,s)-spline interpolation and smoothing.  
Arcangeli, R., de Silanes, M.C.L., Torrens, J.J. Numer. Math, 2007.



# Convergence

The MAP estimate: the conclusion

## Conclusions

The MAP estimate converges to the solution quickly in the sense that:

$$\left| \hat{y}(t) - y(0) - \int_0^t f(\hat{y}(\tau)) d\tau \right| \sim h^\nu. \quad (21)$$

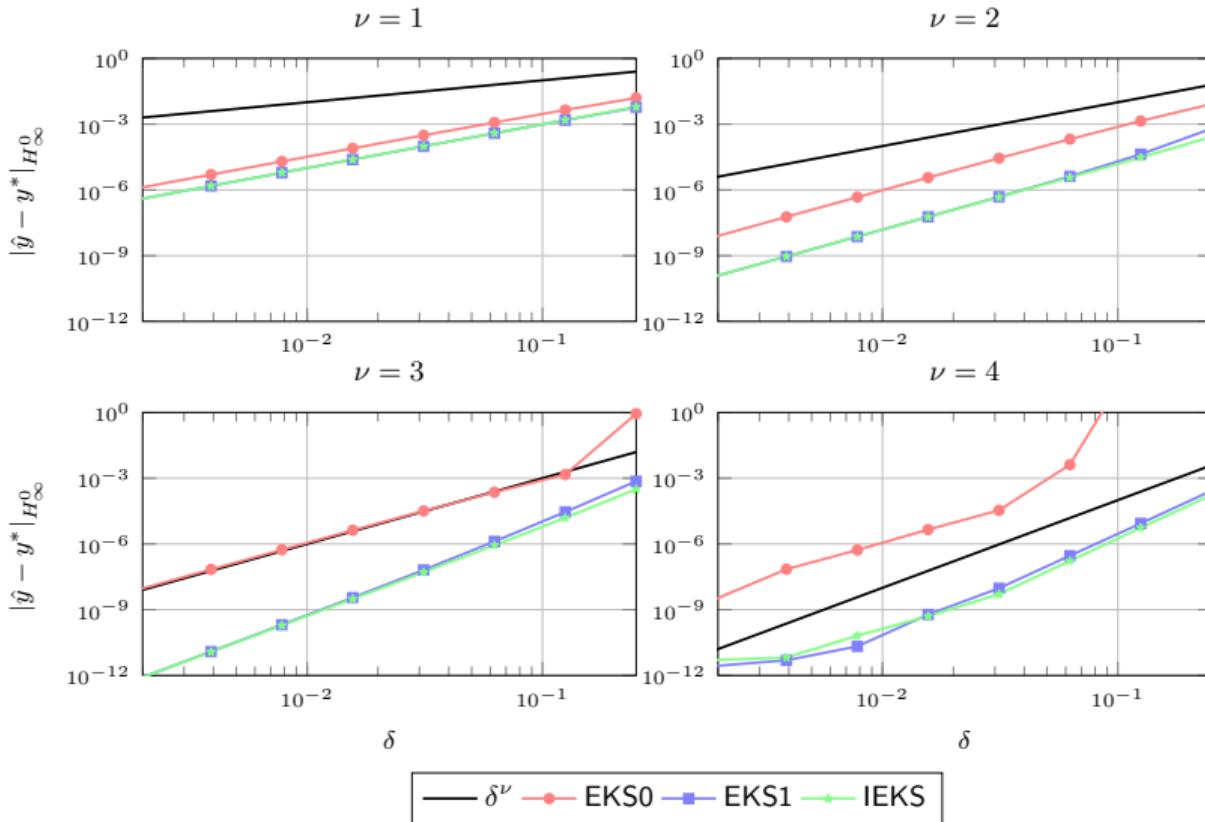
Error estimates may be obtained with Gronwall's inequality:

$$\left| \hat{y}(t) - y^*(t) \right| \sim h^\nu. \quad (22)$$



# Convergence

line goes down





# Convergence

The MAP estimate: some caveats

Some notes:

- The MAP estimate is an idealised object in general (non-convex problem).
- The rates only hold "eventually".
- The minimum norm property has some funky effects:

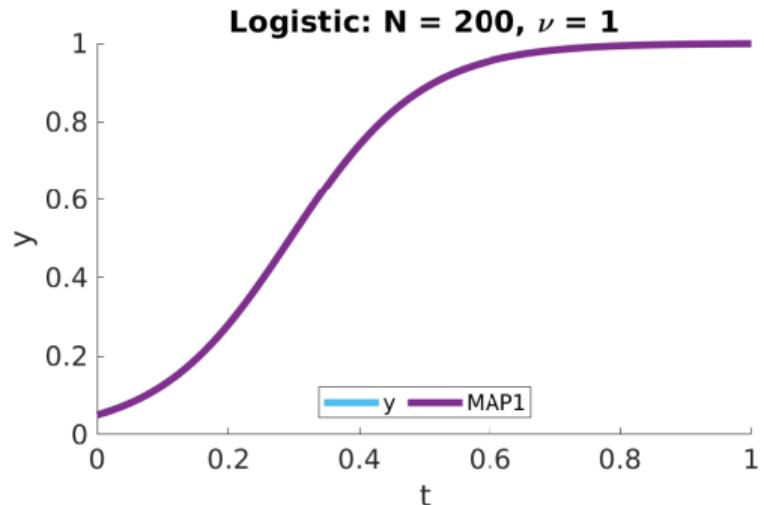
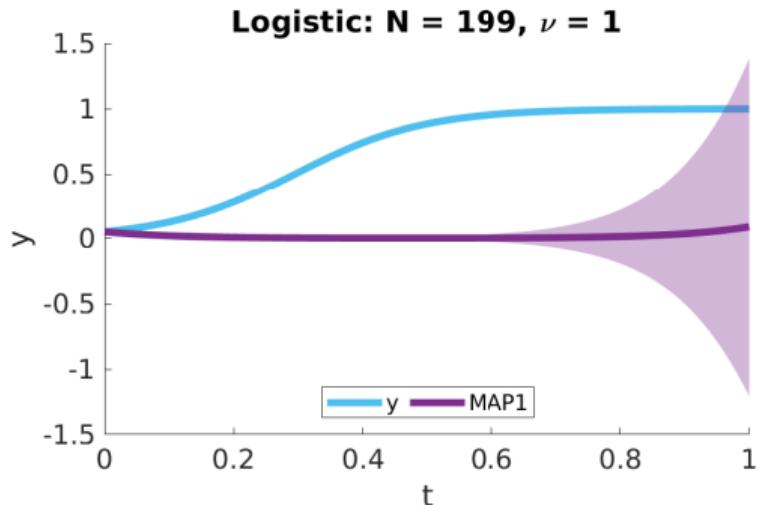
$$\dot{y}^*(t) = ry^*(t)(1 - y^*(t)), \quad y^*(0) = \varepsilon. \quad (23)$$

**There is a function close to zero, which interpolates well.**



# Convergence

The MAP estimate: Eventually can take a while and then happen suddenly





- Estimates can asymptotically die, even if the solution exploded.
- Small functions are premiered, perhaps too much?

RKHS for large time horizons:

$$\|y\|_{\text{RKHS}}^2 = \int_0^\infty \left| \left( y^{(\nu+1)}(t) - \sum_{m=0}^{\nu} A_m y^{(m)}(t) \right) \right|^2 \sigma^{-2}(t) dt \quad (24)$$

**Use  $\sigma$  to make the RKHS norm of the solution small somehow?**



# Compositon of numeric and measurement uncertainty

Problem setting

Parametric ODE:

$$\dot{y}_\theta^*(t) = f_\theta(y_\theta^*(t)), \quad y^*(0) = y_0(\theta).$$

Data:

$$u(t_n) = Hy(t_n) + v(t_n), \quad v(t_n) \sim \mathcal{N}(0, R_\theta). \quad (25)$$

Likelihood functional:

$$L_D(\theta, \varphi) = \prod_n \mathcal{N}(u(t_n); H\varphi(t_n), R_\theta). \quad (26)$$

Marginal likelihood function:

$$M(\theta) = \int L_D(\theta, \varphi) \delta(\varphi - y_\theta^*) d\varphi. \quad (27)$$



# Compositon of numeric and measurement uncertainty

The Marginal likelihood: the probabilistic numerics approach

Output of probabilistic numerics:

$$\hat{\delta}_N(\varphi; \theta, \kappa) \approx \delta(\varphi - y_\theta^*) \quad (28)$$

Marginal likelihood approximation:<sup>19</sup>

$$\hat{M}(\theta, \kappa) = \int L_D(\theta, \varphi) \hat{\delta}_N(\varphi; \theta, \kappa) d\varphi. \quad (29)$$

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<sup>19</sup>Differentiable likelihoods for fast inversion of 'likelihood-free' dynamical systems. H. Kersting, N. Krämer, M. Schiegg, C. Daniel, M. Tiemann, P. Hennig. ICML, 2020.



# Compositon of numeric and measurement uncertainty

The Marginal likelihood: the probabilistic numerics approach

Gauss–Markov representation of  $\widehat{\delta}_N$ :<sup>20</sup>

$$\widehat{\gamma}(x(t_{1:N}); \theta, \kappa) = \mathcal{N}(x(t_N); \xi_\theta(t_N), \kappa \Lambda(t_N)) \prod_{n=N-1}^1 \mathcal{N}(x(t_n); G_\theta(t_n, t_{n+1})x(t_{n+1}) + \zeta_\theta(t_n), \kappa V_\theta(t_n)) \quad (30a)$$

$$\zeta_\theta(t_n) = \mu(t_n) - G_\theta(t_n, t_{n+1})\mu(t_{n+1}^-). \quad (30b)$$

Gauss–Markov regression but backwards:

$$x(t_n) | x(t_{n+1}) \sim \mathcal{N}(x(t_n); G_\theta(t_n, t_{n+1})x(t_{n+1}) + \zeta_\theta(t_n), \kappa V_\theta(t_n)), \quad (31a)$$

$$u(t_n) | x(t_n) \sim \mathcal{N}(Hx(t_n), R_\theta) \quad (31b)$$

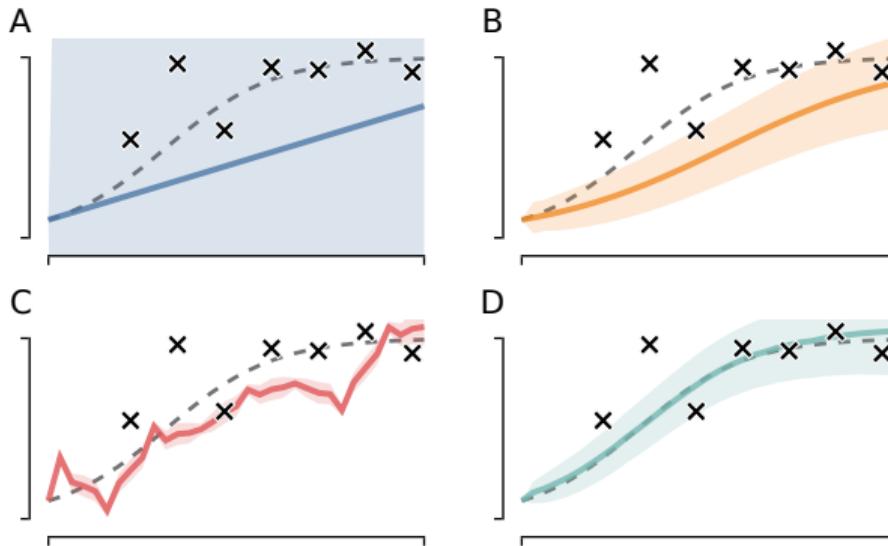
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<sup>20</sup>Fenrir: Physics-Enhanced Regression for Initial Value Problems F. Tronarp, N. Bosch, P. Hennig. ICML, 2022.



# Composition of numeric and measurement uncertainty

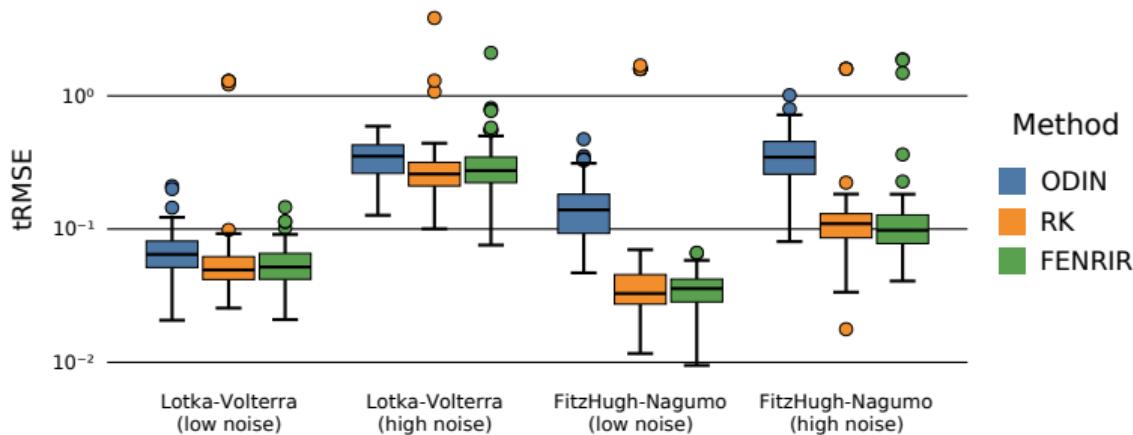
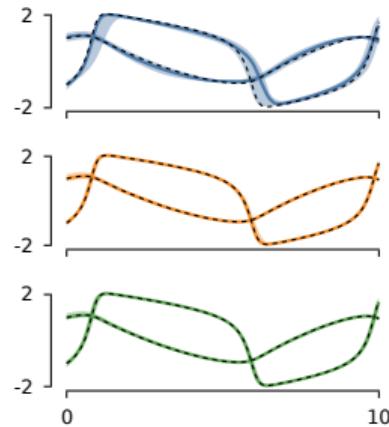
Pictoral numerics





# Composition of numeric and measurement uncertainty

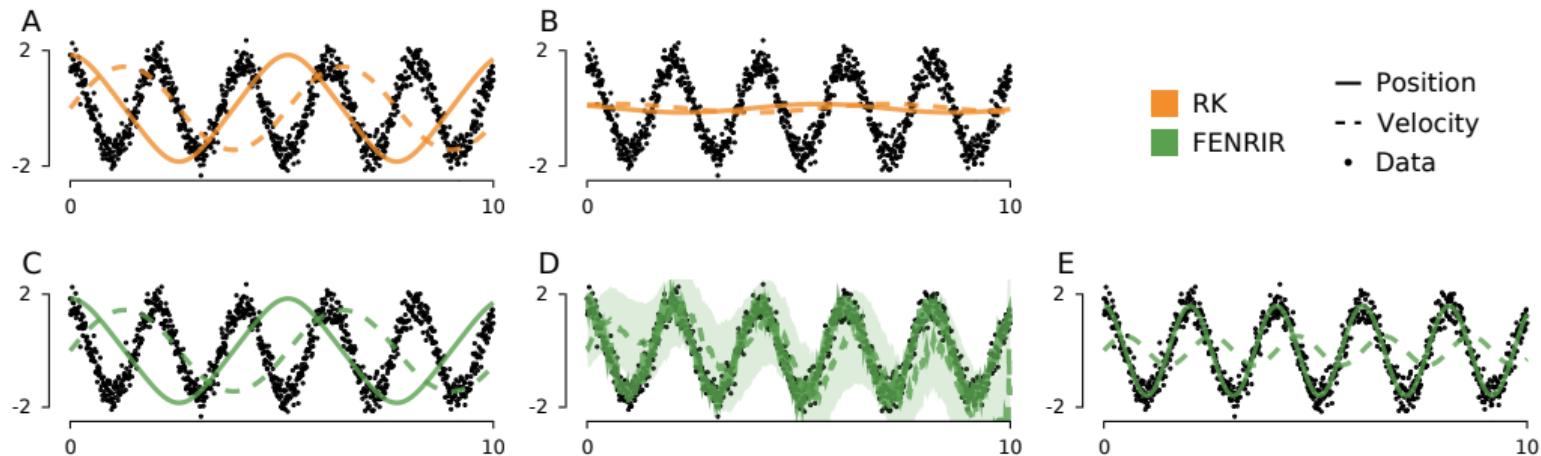
Some benchmarking





# Composition of numeric and measurement uncertainty

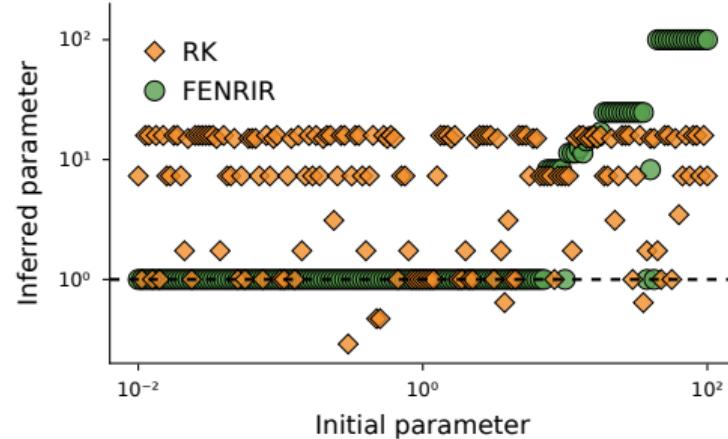
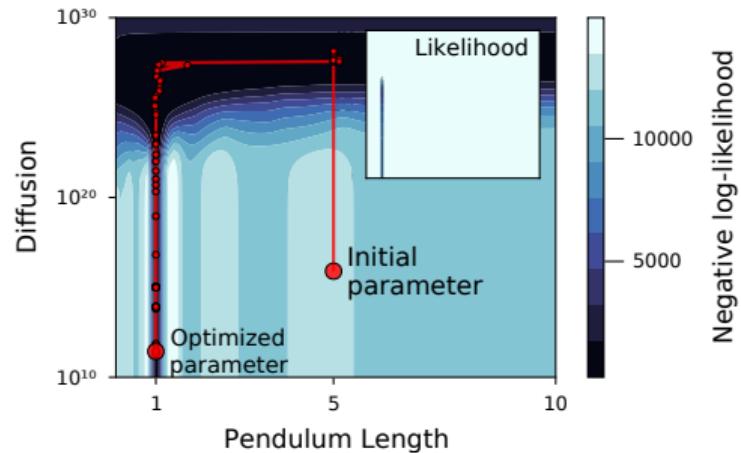
The benefits of likelihood smoothing and uncertainty quantification?





# Composition of numeric and measurement uncertainty

The benefits of likelihood smoothing and uncertainty quantification?





Bells and whistles:

- Numerically stable implementation of probabilistic solvers.<sup>21</sup>
- Solvers for boundary value problems.<sup>22</sup>
- Augmenting measurement model to handle known constraints (e.g. energy conservation).<sup>23</sup>

Software if you care to try:

- Python (ProbNum): <https://probnum.readthedocs.io/en/latest/><sup>24</sup>
- Julia (ProbNumDiffEq.jl): <https://github.com/nathanaelbosch/ProbNumDiffEq.jl>

**Probabilistic ODE solvers are becoming mature in terms of theory, algorithms, and software - BUT!**

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<sup>21</sup>Stable implementation of probabilistic ODE solvers. N Krämer, P. Hennig. arXiv:2012.10106, 2020.

<sup>22</sup>Linear-Time Probabilistic Solutions of Boundary Value Problems. N. Krämer, P. Hennig. Neurips, 2021.

<sup>23</sup>Pick-and-mix information operators for probabilistic ODE solvers. N. Bosch, F. Tronarp, P. Hennig. AISTATS, 2022.

<sup>24</sup>J. Wenger, N. Krämer, M. Pförtner, J. Schmidt, N. Bosch, N. Effenberger, J. Zenn, A. Gessner, T. Karvonen, F.-X. Briol, M. Mahsereci, P. Hennig. ProbNum: Probabilistic Numerics in Python. arXiv:2112.02100, 2021.